

# Econographics\*

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## Abstract

We study the pattern of correlations across a large number of behavioral regularities, with the goal of creating an empirical basis for more comprehensive theories of decision-making. We elicit 21 behaviors using an incentivized survey on a representative sample ( $n = 1,000$ ) of the U.S. population. Our data show a clear and relatively simple structure underlying the correlations between these measures. Using principal components analysis, we reduce the 21 variables to six components corresponding to clear clusters of high correlations. We examine the relationship between these components, cognitive ability, and demographics. Common extant theories are not compatible with all the patterns in our data.

JEL Classifications: C90, D64, D81, D90, D91

Keywords: Econographics, Reciprocity, Altruism, Trust, Costly Third-Party Punishment, Inequality Aversion, Risk Aversion, Common-Ratio Effect, Endowment Effect, WTA, WTP, Ambiguity Aversion, Compound Lottery Aversion, Discounting, Overconfidence, Cognitive Ability, Demographics

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# 1 Introduction

Decades of research in economics and psychology has identified a large number of behavioral regularities—specific patterns of behavior present in the choices of a large fraction of decision-makers—that run counter to the standard model of economic decision-making. This has led to an enormous amount of research aimed at understanding each of these behaviors. However, significantly less work has gone into linking these regularities with each other, either theoretically or empirically. Instead, most regularities have been studied in isolation, with specific models developed for each one. This has led to concerns about model proliferation in behavioral economics.<sup>1</sup> As Fudenberg, (2006, p. 689) notes, “[B]efore behavioral theory can be integrated into mainstream economics, the many assumptions that underlie its various models should eventually be reduced to the implications of a smaller set of more primitive assumptions.”

In this paper, we study the pattern of correlations across a large number of behavioral regularities, with the goal of creating an empirical basis for more comprehensive theories of decision-making. We use an incentivized survey to elicit 21 behaviors from a representative sample of the U.S. population ( $n = 1,000$ ). These *econographics*—a neologism describing measures of trait-like behaviors related to economic decision-making—cover broad areas of social preferences (8 measures), attitudes towards risk and uncertainty (9 measures), overconfidence (3 measures), and time preferences (1 measure). Whenever possible, our elicitations are incentivized: the compensation participants receive depends on their choices. We also include two measures of cognitive abilities and several demographic variables. Moreover, we took steps to limit the effects of measurement error by eliciting many of our measures twice and using the Obviously Related Instrumental Variables (ORIV) technique (Gillen et al., 2019).

Overall, our main finding is that there is a clear and relatively simple structure underlying our data. We can summarize 21 econographics with *six* components, as shown in Section 4, and summarized in Table 1. These components can easily be seen in the correlational structure of the data—for example, the correlation between the measures in the Generosity component range from 0.34–0.86. Two of these components underlie social preferences and

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<sup>1</sup>See, for example, Fudenberg (2006); Levine (2012); Köszegi (2014); Bernheim (2016); and Chang and Ghisellini (2018). There are many important defenses of the wide range of models, including that they may be the most accurate representation of behavior, see, for example, Conlisk (1996); Kahneman (2003); DellaVigna (2009); Chetty (2015); and Thaler (2016).

Table 1: Twenty-one econographics can be summarized with six components.

<i>Components</i>	<i>Social Components</i>			<i>Risk Components</i>		
	<i>Generosity</i>	<i>Punishment (Impulsivity)</i>	<i>Inequality Aversion/WTP</i>	<i>WTA</i>	<i>Uncertainty</i>	<i>Overconfidence</i>
Reciprocity: High		Anti-social Punishment	Dislike Having More	WTP	WTA	Overestimation
Reciprocity: Low		Pro-social Punishment	Dislike Having Less	Risk Aversion: CR Certain	Risk Aversion: Gains	Overplacement
Altruism		(Patience)	Risk Aversion: CR Lottery	Risk Aversion: Losses	Risk Aversion: Losses	Overprecision
Trust			Components combine in joint analysis	Risk Aversion: Gain/Loss		

Notes: Simplified summary of patterns shown in Table 7. The names of econographics correspond with those in the literature. For definitions of specific econographics, see Section 2. As noted in the text, time preferences load weakly on the Punishment component, slightly changing its meaning: this is indicated by the parentheses around that measure, and the alternative name of the component.

beliefs about others, two underlie risk and uncertainty. One component has explanatory power in both the domains of social and risk preferences. A final component underlies overconfidence measures. Time preferences are spread across a number of components, but load most heavily on the Punishment component of social preferences. Each econographic, except for time preferences, loads heavily on only one of the six components. As we detail below, existing models predict some of the relationships we observe, but, to our knowledge, none match the overall pattern, suggesting the need for new theories.

**Approach and Limitations.** We focus on a specific set of measures and study their empirical relationships in order to create an empirical basis that can aid the development of comprehensive theories of decision-making. Our approach is complementary to others in the literature. One standard approach is to focus on specific theoretical links (for example, Chakraborty et al., 2020, which links time and risk preferences). Others start with a specific theoretical mechanism, and then study—theoretically and, in some cases, also empirically— which behaviors it can generate. This is the case of popular modern research programs on rational inattention (Woodford, 2020; Frydman and Jin, 2022; Khaw et al., 2021; Stevens, 2020), salience (Bordalo et al., 2012, 2013), limited strategic thinking (Farhi and Werning, 2019; García-Schmidt and Woodford, 2019), incomplete preferences (Masatlioglu and Ok, 2014; Cerreia-Vioglio et al., 2015), preference imprecision (Butler and Loomes, 2007, 2011), and cognitive uncertainty (Enke and Graeber, 2019). While these important programs are complementary to our approach, our investigation is primarily empirical and does not seek to relate a large number of behaviors to a single underlying mechanism. Rather, it aims to construct a basis for theoretical modeling by exploring the empirical structure of our data, which will turn out to be more complex than what can be explained by a single mechanism.

Examining *all* links between econographics, as we do—including many not studied by theories that seek to link different behaviors—has a number of benefits, especially when compared with the standard approach of theorizing about a particular connection between two behaviors and testing it. First, as correlations are not transitive, identifying underlying structures requires observing all links (or lack thereof) between behaviors.<sup>2</sup> Second, once

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<sup>2</sup>For example, suppose a theory predicts a relationship between behaviors  $A$  and  $B$  that have correlation 0.5. Another theory connects behaviors  $B$  and  $C$ ; a different study finds a correlation of 0.5. Even if there is no theoretically-predicted relationship between  $A$  and  $C$ , their correlation—which can be anywhere between 0 and 1—is important to understanding the true structure. If the correlation is 1, then  $A$  and  $C$  are redundant, so at most one of the two prior theories is correct. If the correlation is 0, then  $B = A + C$ , and

that structure is identified, its components can be examined to see if they are associated with existing constructs—like demographics or cognitive ability. Third, by measuring all behaviors simultaneously in a representative sample, we ensure that the patterns we identify are not due to shifting participant populations between studies.<sup>3</sup>

The empirical structure we uncover could have taken one of three forms, with different implications. First, there may be no discernible structure, suggesting the continued usefulness of examining behaviors in isolation. Second, there may be a structure that matches well with some existing theoretical approach. Third, it could be that a structure exists, but it does not match any extant theory. Our results are somewhere between the second and third possibility: extant theories match *some* of the correlations we observe, but no theory explains the patterns in a specific domain, let alone the overall patterns we observe. This suggests the need for further theoretical exploration, for which we hope we give an empirical basis.

There are inherent constraints in the number and type of behaviors we could elicit on an incentivized survey. This meant we had to leave out many interesting phenomena. While we would like to have been able to capture many other behaviors, making choices was unavoidable. Some measures were infeasible in our context: for example, measuring time preferences using a real-effort task (Augenblick et al., 2015; Cohen et al., 2020) or preferences for competition (Niederle and Vesterlund, 2007), would have taken as much time as approximately a half-dozen of our other elicitation combined. The same holds for other measures, such as inattention, character, and non-cognitive skills, for which standard elicitation are still being developed. Given these constraints, we made a choice and focused on a subset of measures we believe of interest. Future studies can focus on alternative sets of behaviors. Indeed, as we discuss below and in Section 7, some contemporaneous studies already have.

**Analysis and Results.** Describing the 378 correlations between 21 econographic variables, as well as between cognitive and demographic variables, is a daunting task. However, we are aided by the fact that many of the econographics fall into *clusters* of variables—featuring high intra-cluster correlations and low inter-cluster correlations. To summarize them, we make use

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the two theories should be complementary. If the correlation is 0.5, this suggests that an underlying latent factor determines all three behaviors, and the focus should be on developing a theory consistent with it.

<sup>3</sup>Moreover, a correlation that might seem large in isolation (that is, statistically significant), may be quite small in the context of other correlations.

of principal components analysis (PCA), a statistical technique that produces components—linear combinations of variables—that explain as much variation in the underlying behaviors as possible, as discussed in Section 3.<sup>4</sup>

In our analysis, discussed at length in Section 4, we are aided by a fact that was not ex ante obvious: risk and social preferences are largely—but, as we shall see, not completely— independent. Thus, we first study social and risk preferences separately, before combining them together with measures of time preferences and overconfidence.

Our results show there is significant scope for representing our eight measures of social preferences in a more parsimonious way. The behaviors we measure break down into three clusters, as summarized in Table 1. In particular, altruism, trust, and two different types of reciprocity form a cluster. Pro- and anti-social punishment are a second cluster. The third cluster is formed by two different types of inequality aversion—dislike of having more than another person and a dislike of having less. As we discuss below, these results are not in line with any of the existing theories of social preferences.

Risk preferences show a structure that is less parsimonious than standard theory. Our nine measures form three clusters, as summarized in Table 1. Two separate ones contain different measures of risk attitudes towards ordinary lotteries. One cluster is formed by the Willingness to Accept (WTA) for a lottery ticket together with risk aversion for lotteries with gains, losses, and gains and losses as measured by certainty equivalents. The other cluster is formed by the Willingness to Pay (WTP) for a lottery ticket along with risk aversion as measured by *lottery* equivalents. While the members of each cluster are highly related to each other, they are largely independent across clusters, suggesting the presence of two separate and independent forms of risk attitude towards lotteries.<sup>5</sup> These two components have different relationships with other features (for example, age), as shown in Section 5. The third cluster consists of aversion to compound lotteries and ambiguity aversion. These behaviors are highly correlated, consistent with previous studies (Halevy, 2007; Dean and Ortoleva, 2019; Gillen et al., 2019). Our richer data allow us to document that they are

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<sup>4</sup>We also used machine learning techniques that are more sophisticated than PCA, such as clustering, but these did not produce additional insight. See Appendix B.

<sup>5</sup>Note that this finding is distinct from the different “domains” of risk attitudes often discussed in psychology—see Weber and Johnson (2008) for a review—as we examine only the domain of lotteries. Our findings are also distinct from economic studies that document poor correlations between different elicitation of risk attitudes—see Friedman et al. (2014) for a review—as we document a particular pattern of both high and low correlations between measures of risk attitudes.

largely unrelated to other aspects of risk preferences.

Analyzing risk and social preferences together—along with overconfidence and time preferences—results in six components, as summarized in Table 1. Two social components and two risk components are largely unchanged. The three overconfidence measures form a new component.<sup>6</sup> Patience loads most heavily (and negatively) on the punishment component of social preferences. However, one social component—inequality aversion—and one risk component—related to WTP—combine with each other, showing the possibility of further parsimony in the representation. Of course, the generic idea of a connection between risk and inequality aversion has been suggested before. For example, from behind the “veil of ignorance” more inequality creates more risk (Carlsson et al., 2005). Our results indicate that, on the one hand, this has some empirical support (unlike other possible connections that have been suggested). On the other hand, they show how this holds for *only one* of the two components of risk attitudes towards ordinary lotteries (in particular, the one related to WTP), not the other (the one related to WTA).

We also examine the relationships between the components we identify and cognitive abilities and demographics, in Section 5. Four of the six components are correlated with cognitive abilities: higher cognitive ability is associated with a lower propensity to punish and greater patience, lower overconfidence, and higher generosity. Moreover, we document a relationship between cognitive abilities and risk preferences, but only with one of the two components linked with risk aversion—the one joined with inequality aversion. Connecting with demographics, the strongest links are with education and income; moreover, these variables do not seem to be simply proxying for cognitive ability, but have their own individual effect.

**Relation to Theory and Literature.** Finally, we draw out the implications of our data for existing theories, in Section 6, some of which were alluded to above. Broadly speaking, we show that while there is some overlap between our data and common theories of social and risk preferences, no theory explains the overall patterns well, even within a specific domain. For social preferences, the three outcome-based models we consider—altruistic preferences, social welfare, and inequality aversion—predict that many of the measures that

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<sup>6</sup>While an extensive literature has studied different forms of overconfidence and their differences, to our knowledge the finding of a common component, in a large representative survey, is new evidence about this phenomenon (Moore and Dev, 2020).

make up the Generosity component (described in Table 1) should be positively correlated.<sup>7</sup> However, these theories also predict a relationship with inequality aversion, which we do not observe. For risk preferences, models delineating risk and uncertainty do not match our data. Moreover, although reference points are clearly important in explaining the split between risk aversion measures associated with WTA and those with WTP, common theories of reference dependence make additional predictions that are not consistent with our data.

Our study is uniquely suited to our aim of understanding the empirical basis for more parsimonious behavioral models across the risk, social, time, and confidence domains. There is a significant literature, summarized in Section 7, that examines the correlation between two or three behavioral measures, and/or cognitive ability. These more limited sets of relationships are not able to recover all of the the nuanced structure we document here. There are a small number of studies which do measure more behaviors (Burks et al., 2009; Dean and Ortoleva, 2019; Falk et al., 2018; Dohmen et al., 2018; Stango and Zinman, forthcoming). As discussed in Section 7, these studies differ from ours in terms of behaviors examined, incentivization, representativeness, and/or how they deal with measurement error.

## 2 Design and Econographics

Four design decisions followed from our goal of providing an empirical basis for the underlying structure of theories of behavioral decision-making. While the most important was the selection of behaviors to elicit, we return to this discussion after introducing the behaviors we measure.

Our second decision was to study behavioral measures rather than using those measures to estimate model parameters. For example, when measuring risk aversion, we elicit certainty equivalents for lotteries, and use (a linear transform of) them in our analyses, rather than trying to identify a parameter of some utility function (for example a constant relative risk aversion, or CRRA, utility function). This allows our results to be used to inspire theories that connect behaviors without committing to specific functional forms, which are almost surely mis-specified, and for which precise individual-level parameter estimation is difficult

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<sup>7</sup>Outcome-based models are those in which utility is based only on the outcomes received by all players. This is in contrast to models where utility might be based on the set of actions that could have been taken by other players, such as in the fairness model of Rabin (1993).



in a short survey.<sup>8</sup>

Two final design decisions—that our study is incentivized and representative—are easy to justify. Much of the literature we build on comes from laboratory studies in economics, which are almost always incentivized. While there are sometimes good reasons to use non-incentivized measures, these reasons are largely related to feasibility and credibility. These concerns were not substantial in our case, and we were motivated to move past them. In order to make our empirical basis representative of a broad range of people, rather than just specific subgroups, we used a representative sample. These two design decisions drove and constrained a number of implementation details discussed in the next section.

A challenge of using a representative sample is that some participants will be poorly educated (relative to convenience samples of college students). For generalizability, this sampling is a feature, not a bug; but it does mean that elicitation must be simple and designed to have good internal validity for all Americans. As such, many of our behavioral measures are based on the same elicitation technique: indifference elicited using a multiple price list (MPL) method. This was chosen as it allows for more efficient estimation of indifference points than asking individual binary choices—which, for an experiment of our scope, would be infeasible—yet it is seen as easier for participants to understand than incentivized pricing tasks (Cason and Plott, 2014). A training period, with examples and supervised trial sessions, preceded the actual survey.<sup>9</sup> Other common issues that arise in trying to draw representative samples are addressed in Section 3.1. The techniques we used to (statistically) deal with measurement error, detailed in Section 3.2, further helped in recovering valid estimates from these low-education populations.

Appendix A gives implementation details omitted here. The specific question wordings, screenshots, and other details of experimental design can be found in the supplementary materials uploaded with this submission and online at [eriksnowberg.com/wep.html](http://eriksnowberg.com/wep.html).

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<sup>8</sup>Additionally, recovering a parameter of a behavioral model through non-linear transformations of quantities measured with error is problematic. For example, some participants state relatively high, or low, certainty equivalents for lotteries that result in huge (positive and negative) CRRA coefficients. These values dominate correlations, causing them to be more informative about measurement error than behavior. Upon publication, our data will be made publicly available, and researchers interested in particular parametric formulations can use them to test those theories, if they so desire.

<sup>9</sup>We also included dominated options at the endpoints of the MPL scales wherever possible. The undominated options in these rows were pre-selected, following Andreoni and Sprenger (2012a). The software also imposed a single crossing point.

## 2.1 Social Preferences

There are many examples where people’s actions take into account the preferences and beliefs of others, even in non-strategic settings. The motivating factors behind these acts are often given the broad term of social preferences.

Altruism is defined as giving to strangers while expecting nothing in return. In experimental economics, it is usually measured using the dictator game (Forsythe et al., 1994; Falk et al., 2013), in which one participant decides unilaterally how to split money between themselves and another person. Following this literature, we measure **Altruism** as the amount given to another person in the dictator game.

Trust and reciprocity are intertwined. To understand why, imagine that a stranger asks for money for a sure-thing investment. In order to provide money in such a venture you must *trust* that he will give some of the proceeds back to you. The act of giving money back is *reciprocation*, which may depend on how much money you gave to the stranger. We measure these concepts through a standard trust game: one participant (the sender) decides how much of an endowment to send to a second (the receiver). This amount is doubled by the experimenter, and then the receiver decides how much to send back—which is also doubled (Berg et al., 1995). We measure **Trust** as the amount sent by the participant when they are in the role of the sender. The original sender will also take the role of receiver in a different interaction: **Reciprocity: Low** corresponds to the amount sent back when receiving the lowest amount, and **Reciprocity: High** corresponds to the amount sent back when receiving the maximum amount.<sup>10</sup>

People are willing to punish others for what they perceive as bad behavior—even when that behavior does not directly affect them—and punishment is costly. To measure this, we allow participants to observe a trust game in which the sender gives all the money they have, and the receiver returns nothing. We give each participant a stock of points to punish the receiver—that is, to pay a cost to reduce the points the receiver gets: the amount used is **Pro-social Punishment**. Prior studies document that a significant minority of people—

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<sup>10</sup>These two types of reciprocity are sometimes referred to as positive and negative reciprocity. However, negative reciprocity is usually defined as the response to the lowest possible action. In this case, the lowest action is sending no money, to which the receiver cannot respond. Thus, to avoid confusion with the standard usage, we use low and high rather than negative and positive. Note also that each participant’s partner in these interactions differs between when they are the sender and the receiver. Moreover, when the participant plays the role of receiver, we use the strategy method: that is, we elicit the response for every possible amount sent. For more details of implementation, see Appendix A.

the percent possibly depending on culture—also punish the sender (Herrmann et al., 2008). Thus, we also give a separate stock of points that can be used to punish the person who sent all their money. The amount used to punish the sender is **Anti-social Punishment**.

Many people seem uncomfortable with having a different amount (greater or less) than others, a phenomenon known as *inequality aversion* (Fehr and Schmidt, 1999; Charness and Rabin, 2002; Kerschbamer, 2015). **Dislike Having Less** is how much a person is willing to forgo in order to ensure that they will not have less than another person. **Dislike Having More** is how much a person is willing to forgo in order to ensure that they will not have more than another person.

## 2.2 Measures of Risk Attitudes

To measure attitudes towards risk and uncertainty, we elicit the valuation of various prospects. All lotteries involve only two possible payoffs, and most assign 50% probability to each.

Following the standard approach, we identify the behavioral manifestation of risk aversion as valuing a lottery at less than its expected value. Extensive research shows that the patterns of valuation depend on whether a lottery contains positive payoffs, negative payoffs, or both a positive and a negative payoff.<sup>11</sup> Thus, we include three measures of risk aversion: **Risk Aversion: Gains** elicits a participant’s certainty equivalent for a lottery containing non-negative payoffs, **Risk Aversion: Losses** elicits a participant’s certainty equivalent for a lottery with non-positive payoffs, and **Risk Aversion: Gain/Loss** elicits a participant’s certainty equivalent for a lottery with one positive and one negative payoff (Cohen et al., 1987; Holt and Laury, 2002). The difference between the expected value of the lottery and a participant’s value are used in the analysis, so larger numbers indicate more risk aversion.

The endowment effect is the phenomenon that, on average, people value a good more highly if they possess, or are *endowed* with, it. In our implementation, **WTP** is the amount a participant is *willing to pay* for a lottery ticket, and **WTA** is the amount the participant is *willing to accept* for the same ticket when she or he is endowed with it. The difference between WTA and WTP is the **Endowment Effect** (Kahneman et al., 1990). We discuss this as a risk preference because the object being bought and sold is a lottery ticket, but more importantly, because of the patterns revealed in the analysis of Section 4.3. However,

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<sup>11</sup>The most familiar theoretical model of these patterns of valuation is Prospect Theory (Kahneman and Tversky, 1979).

as the pattern of correlations in our data suggest that WTA and WTP are also fundamental behaviors, we primarily examine these rather than the endowment effect.

Risk attitudes often change when one of the available options offers certainty, as demonstrated through the common ratio effect (Allais, 1953). Under Expected Utility, when the winning probabilities of two lotteries are scaled down by a *common* factor, a person’s ranking over those lotteries should not change. However, this is often not the case. To capture this effect, we ask the participant to make two choices, one that measures risk aversion with a certain alternative, and another in which both options are risky. In **Risk Aversion: CR Certain**, we elicit the amount  $b$  such that the participant is indifferent between a certain amount  $a$  and a lottery paying  $b$  with probability  $\alpha$  (and zero otherwise): that is, a lottery equivalent of a sure amount. In **Risk Aversion: CR Lottery**, we elicit the amount  $c$  such that the participant is indifferent between a lottery paying  $a$  with probability  $1/x$  (and zero otherwise) and  $c$  with probability  $\alpha/x$  (and zero otherwise). Under Expected Utility,  $b = c$ ; the **Common Ratio** measure is then  $b - c$  (Dean and Ortoleva, 2019). In keeping with our treatment of the endowment effect, we enter the constituent measures in most analyses.<sup>12</sup>

Ambiguity aversion is a preference (or beliefs that lead to a preference) for prospects with known probabilities over those with unknown probabilities. To measure it, we use an *ambiguous urn* filled with balls of two different colors. One color gives the participant a positive payoff, and the other gives them zero. Participants do not know the proportion of the different color balls in the urn, but are allowed to choose which color gives a positive payoff. They are then asked for their certainty equivalent for a draw from this urn. If participants have a prior over the composition of the urn, they must believe a draw from the urn has a winning probability of at least 50%, yet many participants prefer a 50/50 lottery with known odds. The difference between the certainty equivalent for a draw from the *risky urn*—with a known composition of 50% of each color—and the ambiguous urn is **Ambiguity Aversion**.<sup>13</sup> Similarly, a draw from a risky urn is usually more highly valued than one in which the number of balls is unknown, but drawn from a uniform distribution—that is, a

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<sup>12</sup>We check the robustness of this analysis to adding the single common ratio measure in Appendix Table E.2.

<sup>13</sup>The risk measure here is **Risk Aversion: Urn**, which matches the description above. Empirically, this measure is highly correlated with Risk Aversion: Gains. Note that both Compound Lottery Aversion and Ambiguity Aversion difference out the same quantity. If this quantity is measured with error, this can create a spurious correlation between Compound Lottery Aversion and Ambiguity Aversion. This issue, and our solution, is discussed in Section 3.2.

compound lottery. The difference between the certainty equivalents for a draw from the risky urn and one from a compound urn is **Compound Lottery Aversion** (Halevy, 2007).

## 2.3 Overconfidence

Overconfidence can be divided into three types. **Overestimation** refers to a person’s estimate of her performance on a task (versus her actual performance). **Overplacement** refers to her perceived performance relative to other participants (versus her real relative performance). In order to measure these phenomena, we ask participants to complete two tasks: on one, we ask them to estimate their performance, and the other, to estimate their performance relative to others taking the survey. The difference between these subjective estimations and actual performance, in absolute or percentile terms, give us overestimation and overplacement, respectively (Moore and Healy, 2008).

**Overprecision** refers to a belief that one’s information is more precise than it actually is. We ask participants to estimate a number (such as the year the telephone was invented), and then tell us how close they think they were to the correct answer. To difference out overprecision from justified precision, we regress how close the participant thought they were on a fourth order polynomial of their accuracy (Ortoleva and Snowberg, 2015b,a).<sup>14</sup>

## 2.4 Patience (Time Preferences)

A payoff sometime in the future is generally seen as less valuable than a payoff of the same size today. The value today of a fixed future payoff is **Patience** (Andersen et al., 2008).<sup>15</sup>

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<sup>14</sup>Note that as both overprecision and overplacement refer to the same factual questions, this will create correlated measurement error between them. This can create spurious correlation between the two measures. This issue, and our solution, are discussed in Section 3.2.

<sup>15</sup>In an earlier survey we attempted to measure present bias, but found little or no evidence of it, similar to some recent studies that attempt to elicit present bias using both financial payments and effort (see, for example, Augenblick et al., 2015; Imai et al., 2021). This may be due to the fact that points in our study were not (usually) instantly convertible into consumption. Thus, we did not attempt to measure it here. A general concern about experiments that attempt to measure time preferences is whether the participants trust the experimenter to follow through with payment (Andreoni and Sprenger, 2012a). Using a survey company with established relationships with its panelists seems to have largely mitigated this, as discount rates were quite low ( $\delta$  was quite high).

## 2.5 Choice of Measures

The choice of measures to include was driven largely by a desire to focus on “the basics,” and time constraints inherent in the study design. This general logic still left room for many subjective choices. In this subsection, we try to lay out how this logic was applied, and how subjective considerations ultimately led to the set of measures we focused on.

Most of our measures are simple deviations from the standard neoclassical model of a selfish, expected-utility maximizer. Behavior under that model is relatively simple: Choices involving risk and uncertainty are driven by the curvature of the utility function. Adding time adds a discount rate. Choices involving other individuals are no different than choices involving the market: all result in maximizing the welfare of the individual, with no regard for the welfare of others. Finally, complex choices are processed no differently from simple ones. In the domain of risk, deviations from this model allow for distortions in the way these choices are perceived, including reference points and violations of independence. In the domain of social choice, deviations allow for caring about the welfare of others. Deviations in the domain of cognitive limitations include biased cognitive processing, how those biases affect strategic interactions, and so on.

We focused mainly on preferences in the risk and social domains. These areas have generated enormous interest both within and beyond behavioral economics. Arguably, they constitute the areas where behavioral economic models have been most extensively applied in other fields (for example, finance and development economics). Each area has generated multiple empirical regularities that are inconsistent with the standard model of economic decision-making, leading to model proliferation. As different modeling approaches suggest different empirical relations between these regularities, it is particularly useful to establish the correlational structure between these behaviors to guide future modeling work.

Within each domain, the measures we included were guided by a combination of theory and importance to the literature. In the case of risk, we first focused on the two classic violations of independence: common ratio and ambiguity aversion measures. Both have been central to generating huge literatures on non-expected utility theory. Moreover, both have been explained as manifestations of probability weighting (for example, Quiggin, 1982; Segal, 1987), as has aversion to compound lotteries (Segal, 1990). Probability weighting has also been advanced as a possible explanation for small-stakes risk aversion.

We then focus on the second major behavioral force in choice under risk: reference

dependence. Since the seminal work of Kahneman and Tversky (1979), the ideas of reference-dependent risk attitudes and loss aversion have been central to understanding risky choice. The key behavioral patterns associated with reference dependence are differing risk attitudes in the gain and loss domain, and an increase in risk aversion for lotteries that include both gains and losses, thus motivating our three risk aversion measures. At the same time, the idea of loss aversion has been used to explain the endowment effect (see, for example, Kahneman et al., 1990; Kőszegi and Rabin, 2007). This motivated us to also include a measure of the endowment effect in our study.

We would argue that our 9 measures of risk attitude span the behaviors that have been most influential in guiding the development of models of risky choice, and are the most obvious manifestations of the three key theoretical constructs in the literature—ambiguity attitudes, non-linear probability weighting, and reference dependence.

In the case of social preferences, we selected a set of measures that both reflect standard experiments and connect to most widespread theories of social preferences in a parsimonious way. In virtually all models of social preferences, people account for the other’s well-being. This is most directly captured by our measures of altruism and trust, which are both associated with classical measures in their own right. Common approaches, most prominently (Fehr and Schmidt, 1999), add the role of inequality aversion, with a specific distinction between preferences over advantageous and disadvantageous inequality. Other theories point towards the importance of reciprocity and different types of punishment. Importantly, different theories make different predictions about the distribution and correlation of these behaviors.

We made less of an attempt to cover behaviors thought to be driven by biases or mistakes, largely due to time constraints on the survey. Moreover, many of these judgment biases are carefully considered in (Stango and Zinman, forthcoming), which can be seen as a companion to our more intensive focus on preferences. However, we did have space for a couple of such behaviors, and chose to focus on overconfidence (Moore and Healy, 2008) and level-K thinking (Nagel, 1995). Overconfidence was motivated by a longer-term interest in the relationship between economic and political behavior, and prior results suggest that overconfidence is an important link between the two (Ortoleva and Snowberg, 2015a,b). Level-K thinking has long been a subject of interest for one of our co-authors, although our elicitations did not seem to produce reliable results, see Appendix Figure E.1.

Naturally, there are other areas arguably of equal interest to behavioral economists that we have ruled out on various grounds. While there are many interesting behavioral phenomena and traits in strategic settings—for example, quantal response and competitiveness—the types of questions needed to measure these behaviors comprehensively would be too time-consuming for our survey. For similar reasons, we were only able to measure time preferences using money, rather than, say, time dated real-effort tasks. Given the well known issues with this approach, we restricted ourselves to a single time preference measure. In the interests of parsimony, we measured reference dependence within the context of risk preferences rather than separately, and did not measure violations of the independence of irrelevant alternatives, such as the compromise or asymmetric dominance effect. Overall, time constraints in the survey forced us to make choices and to leave out many measures that would be interesting to explore in future work.

### 3 Implementation and Analysis

This section describes our representative, incentivized survey, and the statistical techniques used to eliminate the attenuating effects of measurement error.

#### 3.1 Survey Implementation

Administering an incentivized survey to a representative population presents challenges not normally dealt with in lab environments. To surmount these challenges, we partnered with YouGov, a worldwide leader in online surveys serving the public, businesses, and governments.<sup>16</sup> Our study was given to a representative sample of 1,000 U.S. adults between March 30 and April 14, 2016. We consulted extensively with YouGov on our study design to utilize its expertise in survey design and implementation.

Constructing a representative sample is difficult given variation in response rates. In order to do so, most modern surveys weight on demographics. YouGov supplements this with its own panel of participants. It continually recruits new people, especially from difficult-to-reach and low-socio-economic-status groups. To generate a representative sample, it randomly draws people from various Census Bureau products, and matches them on observables to members of its panel. Oversampling and differential response rates lead to the over- and

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<sup>16</sup>Companies that use in-person or phone surveys, such as Gallup, were unable to administer incentives.



under-representation of certain populations. Thus, YouGov provides sample weights to recover estimates that would be obtained from a fully representative sample. According to Pew Research, YouGov’s sampling and weighting procedure yields more representative samples than traditional probability (that is, random) sampling methods, including Pew’s own probability sample (Pew Research Center, 2016, YouGov is Sample I). We use these weights throughout the paper.<sup>17</sup>

Incentivized questions pose additional challenges: stakes, and whether the experimenter is seen as credible in making future payments or running randomizations as specified. Two randomly selected questions were chosen for payment.<sup>18</sup> To enhance the credibility of our study, we took advantage of YouGov’s relationship with its panel, and restricted the sample to those who had already been paid (in cash or prizes) for their participation in surveys.

All outcomes to incentivized questions were expressed in points. This is an internal YouGov currency used to pay participants. It can be converted to U.S. dollars, or prizes, using the approximate rate of \$0.001 per point.<sup>19</sup> The average payment to participants was around 9,000 points (or \$9). The survey took participants between 45 minutes and an hour. This compensation level is quite high for an internet survey, and represents a rate of pay approximately three times the average for similar surveys.

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<sup>17</sup>As economists rarely run their own surveys in representative populations, it is worth explaining how the survey research literature uses the term “representative.” With few exceptions—censuses, and samples in rural areas of developing countries based on a census—representative samples are representative on observables, not on unobservables. While random samples have the potential to be representative on both observable and unobservables, low response rates render these samples less representative on both observables and expressed preferences, as the Pew study documents. Commonly used representative surveys in economics, such as the Current Population Survey, use weighting to account for non-response. The CPS also uses imputation to adjust for item non-response, which is not present in our survey (see [www.census.gov/programs-surveys/cps/technical-documentation/methodology/imputation-of-unreported-data-items.html](http://www.census.gov/programs-surveys/cps/technical-documentation/methodology/imputation-of-unreported-data-items.html)).

<sup>18</sup>We chose to pay for two randomly selected questions to increase the stakes while making fewer participants upset about their payoffs. Paying for two questions instead of one may theoretically induce some wealth effects, but these are known to be negligible, especially in an experiment such as ours (Charness et al., 2016). Paying for randomly selected questions is incentive compatible under Expected Utility, but not necessarily under more general risk preferences, where it is known that no such mechanism may exist (Karni and Safra, 1987; Azrieli et al., 2018). A growing literature suggests this theoretical concern may not be empirically important (Beattie and Loomes, 1997; Cubitt et al., 1998; Hey and Lee, 2005; Kurata et al., 2009), but there are some exceptions (Freeman et al., 2019).

<sup>19</sup>The conversion from points to awards can only be done at specific point values, which leads to a slightly convex payoff schedule. This is of little concern here, as these cash-out amounts are further apart than the maximum payoff from this survey.

## 3.2 Measurement Error

Measurement error causes a downward bias in correlations. This would make it more difficult to pick out the clusters of inter-related measures. Moreover, this downward bias would also make it difficult to know which behaviors are actually *not* related. To circumvent this issue, we use the ORIV technique of Gillen et al. (2019), which uses the instrumental variables approach to errors-in-variables to produce efficient, consistent estimates. At the heart of this method is duplicate observations of variables—multiple elicitations that are similar, but not exactly the same—that are likely to have orthogonal measurement error. This technique takes duplicate measures and estimates a stacked regression in which each measure of one behavior is used as both an independent variable and an instrument. This is done twice, once for each measure of the other behavior, effectively averaging across all four specifications in the stacked regression model.<sup>20</sup>

Our setting requires us to deal with an additional issue: constructed variables will often have correlated measurement error due to the nature of their construction. For example, the Compound Lottery Aversion and Ambiguity Aversion measures are both constructed by taking some behavior (the certainty equivalent for a compound or ambiguous urn), and subtracting off the same quantity measured with error (certainty equivalent of a risky urn). This leads to correlation in the measurement error of Ambiguity and Compound Lottery Aversion. To avoid spurious correlations, we make use of the fact that we have two observations for each measure, and modify the ORIV procedure so that the measurement error in the instrument is uncorrelated with the measurement error in the left-hand-side variable.<sup>21</sup> We use this formulation for all sets of variables constructed from two elicitations: Compound Lottery and Ambiguity Aversion, and Overplacement and Overprecision.

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<sup>20</sup>Our approach to the design of the multiple measures follows directly from the recommendations in Gillen et al. (2019). When we do not have an available duplicate for one of the measures, we use an approximation of ORIV detailed in Footnote 31 of that paper.

<sup>21</sup>Formally,  $X^i$  and  $Y^i$ ,  $i \in \{1, 2\}$  are two elicitations of  $X^*$  and  $Y^*$  measured with error. For a given  $i$ ,  $X^i$  and  $Y^i$  are constructed by subtracting the same quantity (measured with error) from two different quantities. Thus, the measurement error in  $X^i$  and  $Y^i$  are correlated for a given  $i$ , but the measurement error in  $X^i$  and  $Y^j$ ,  $i \neq j$  are not. The stacked regression in ORIV is then modified to become:

$$\begin{pmatrix} Y^a \\ Y^b \end{pmatrix} = \begin{pmatrix} \alpha_1 \\ \alpha_2 \end{pmatrix} + \beta \begin{pmatrix} X^a \\ X^b \end{pmatrix} + \eta, \quad \text{with instruments } W = \begin{pmatrix} X^b & 0_N \\ 0_N & X^a \end{pmatrix}.$$

### 3.3 Multiple Hypothesis Testing

This paper displays a large number of correlations and standard errors. There are no theoretical predictions for most of the correlations we examine. Thus, we omit any description of statistical testing or significance from our tables. However, if one were interested in null-hypothesis statistical testing (NHST), the appropriate critical value for significance at the 5% level is between 1.96 for a single hypothesis test, and 3.82 using a Bonferroni (Dunn, 1958, 1961) correction for all 378 correlations underlying Sections 4 or 5.

### 3.4 Principal Component Analysis

We examine correlation matrices directly, and use principal component analysis (PCA) to summarize them. As PCA is an analysis of a correlation matrix, and our correlation matrices are corrected for measurement error, so too are the results of PCA. The aim of PCA is to extract the  $m$  components most useful for explaining  $n > m$  variables. Components are linear combinations of the variables. The first component is constructed to capture the highest possible fraction of variance in the data (subject to the constraint that the linear weights sum to one), the second to capture the highest fraction of the remaining variance, conditional on being orthogonal to the first component, and so on.<sup>22</sup>

Once components are identified, the key question is, “How many are necessary to provide a good description of the underlying data?” Heuristically, we want to retain components only when the marginal explanatory power is high. In order to determine the number of components to retain, we use an approach which captures this intuition: parallel analysis. Parallel analysis creates many random datasets with the same numbers of observations and variables as the original data. The average eigenvalues of the resulting correlation matrices are then computed. Components are kept as long as their associated eigenvalues are greater in the actual data than the average in the randomly generated data.<sup>23</sup>

The retained components help to understand the relationship between the original variables in the dataset. The correlation between a component and a variable is called the variable’s *loading* on that component. Variables that load heavily on the same component

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<sup>22</sup>See Abdi and Williams (2010) for an introduction. We also tried more sophisticated machine learning techniques, such as clustering, but these produced no additional insight. See Appendix B.

<sup>23</sup>The value of additional components is graphically represented by a “scree” plot. This shows the eigenvalues of the correlation matrix of the underlying data, and compares this with the average value of the eigenvalues produced by parallel analysis. The scree plots for our analysis can be found in Appendix C.

are highly related. In order to facilitate interpretation, retained components can be rotated relative to the data. Following standard practice, we rotate the resulting components using the Varimax rotation (Furr, 2017, p. 92). This rotates the basis identified from the retained components to maximize the variance of the squared loadings, easing interpretation of the components. However, as we will see, the patterns in the components largely line up with apparent patterns in the correlation matrices, making interpretation straightforward.

In Section 5, we analyze whether the components we identify correlate with cognitive and demographic measures. As we have two elicitations of most variables, to compute the components we multiply the weights from the PCA by the average of the two elicitations.

## 4 Relationships between Econographics

The next two sections attempt to explain, as succinctly as possible, the relationships between 21 econographics variables.<sup>24</sup> We begin by examining each econographic separately. Next, we study the relationships between econographics through a visual inspection of correlation matrices, followed by principal components analyses to verify the observed patterns. This leads to our central finding that the 21 econographics are well summarized by six principal components. In the next section, we examine the relationship between these principal components and cognitive abilities and demographics. As noted in the introduction, there are many ways one might summarize these 378 relationships. Our approach is driven by the desire to create an empirical basis for an underlying structure of more comprehensive theories of behavioral decision-making, and by the hope that this is a relatively straightforward way to do so.

We describe the relationships between econographics in three steps. First, we examine social preferences, then risk preferences, and then combine social and risk preferences together with overconfidence measures and patience. This is done for simplicity, and because each of the first two types of preferences may be of independent interest. Moreover, as our results are largely driven by clusters of correlations, these clusters will not disappear when additional measures are added to the analysis (although they may be augmented).

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<sup>24</sup>We have chosen to present our results first, then relate them to previous empirical results afterward. Given the number of results here, it is difficult to preview them all, relate them to the literature, and then go back and describe them all in detail.

## 4.1 Summary Statistics

The summary statistics in Table 2 show that behavior in our data is consistent with standard findings in the laboratory. In addition to summary statistics, we also show the percent of participants whose responses are in the “expected” direction.<sup>25</sup> Surveying the information in this column, the majority of participants are risk averse or risk neutral over gains, risk loving over losses, inequality averse, exhibit an endowment effect and the common ratio effect, are ambiguity averse, have a negative reaction to compound lotteries, and are overconfident. (A majority of participants are not all of these things simultaneously.) Note that although patience in this table is represented as a discount rate, in the correlation analysis we code the variable as discussed in Section 2.4. Either coding gives the same (directional) interpretation to correlations. However, the coding in Section 2.4 is linear in a participant’s answer, allowing for the measurement error correction discussed in Section 3.2.

Our data exhibit fairly standard levels of noise. As discussed in Gillen et al. (2019) and Snowberg and Yariv (2021) the correlation of duplicate measures—subtracted from one—gives the level of noise in a particular elicitation. Gillen et al. (2019) report correlations of around 0.65 between duplicate measures, using data from Caltech undergraduates. In most cases, the correlations we observe—in the final column of Table 2—are somewhat higher. This implies that our data is *less* noisy than similar data obtained from Caltech undergraduates. The exceptions are the overconfidence measures, which are noisier than the rest.<sup>26</sup> When there is no correlation listed, we only have one elicitation of that behavior.

## 4.2 Links Between Social Preferences

There is ample opportunity to create a more parsimonious representation of the social preferences we measure: altruism, trust, anti- and pro-social punishment, and distributional preferences. These measures fall into three clusters shown in correlation Table 3: one formed by the two measures of reciprocity, and altruism and trust; a second formed by the willingness to punish pro- and anti-social behavior; and a third formed by our two measures

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<sup>25</sup>One could instead code people as having “expected” or “unexpected” behavior on a dimension, or define a point on each dimension that is “neutral” (such as risk-neutral) and then measuring “biases” in either direction from that point. These approaches are closer to that of Stango and Zinman (forthcoming). As our focus is the behavior itself, rather than on how behavior compares to predictions of a model, we examine behaviors linearly in the scale they are measured.

<sup>26</sup>Gillen et al. (2019) do not report data from overconfidence measures.

Table 2: Summary Statistics of Econographics Measures

Variable	Description / Unit	Mean Value	Standard Deviation	% Expected Direction	Corr. btw. Duplicates
Reciprocity: Low	% of Possible Points Returned	0.42	0.23		0.80
Reciprocity: High	% of Possible Points Returned	0.39	0.22		0.96
Trust	% of Possible Points Sent	0.45	0.26		
Altruism	% of Possible Points Sent	0.41	0.27		
Anti-social Punishment	% of Possible Points Used	0.21	0.35		
Pro-social Punishment	% of Possible Points Used	0.50	0.42		
Dislike Having More	% of Income Forgone for Equal Split	0.08	0.45	78%	0.74
Dislike Having Less	% of Income Forgone for Equal Split	0.02	0.37	58%	0.69
Risk Aversion: Gains	$(EV - CE)/EV$	-0.04	0.49	50%	0.64
Risk Aversion: Losses	$(EV - CE)/EV$	-0.29	0.52	73%	0.69
Risk Aversion: Gain/Loss	$(EV - CE)/EV$	-0.04	0.55		0.71
WTA	% of Expected Value	0.91	0.41		0.70
WTP	% of Expected Value	0.65	0.36		0.75
Endowment Effect	% of Expected Value	0.26	0.57	74%	0.75
Risk Aversion: CR Certain	1 - EV of LE as % of Sure Amount	-0.32	0.44	71%	0.76
Risk Aversion: CR Lottery	1 - EV of LE as % of EV of Lottery	-0.35	0.44	73%	0.70
Common Ratio	RA: CR Certain - RA: CR Lottery	-0.03	0.52	62%	0.62
Ambiguity Aversion	(Risky CE - Ambiguous CE)/EV	0.07	0.55	71%	0.64
Compound Aversion	(Risky CE - Compound CE)/EV	0.03	0.56	68%	0.57
Overestimation	Perceived - Real # Correct (of 3)	0.69	1.18	87%	0.37
Overplacement	Perceived - Real Percentile	6.4	39	60%	0.30
Overprecision	Standardized Subjective Precision	-0.02	1.00		0.42
Patience	Monthly Discount Rate	0.87	0.23		0.78

Notes: EV = Expected Value, CE = Certainty Equivalent, LE = Lottery Equivalent. % Expected Direction refers to the percent of the participants (weighted) that give an answer in the direction expected given the current literature: risk averse or risk neutral for most risk questions; risk loving or risk neutral for risk aversion over losses; equality seeking for distributional preferences; endowment effect greater than 0; overweighting small probabilities for common ratio; and ambiguity / compound averse or neutral. When there are two measures of a quantity, those measures are normalized and stacked, so the sample statistics are drawn from 2,000 observations from 1,000 people.

Table 3: ORIV Correlations of Social Measures

	Reciprocity: Low	Reciprocity: High	Altruism	Trust	Anti-social Punishment	Pro-social Punishment	Dislike Having More	Dislike Having Less
Reciprocity: Low		0.86 (.027)	0.34 (.063)	0.49 (.052)	-0.03 (.062)	0.06 (.058)	0.16 (.053)	-0.07 (.054)
Reciprocity: High	0.86 (.027)		0.34 (.053)	0.49 (.041)	-0.03 (.053)	0.08 (.050)	0.22 (.048)	-0.05 (.054)
Altruism	0.34 (.063)	0.34 (.053)		0.60 (.040)	0.05 (.049)	0.03 (.047)	0.25 (.057)	-0.04 (.061)
Trust	0.49 (.052)	0.49 (.041)	0.60 (.040)		-0.05 (.041)	0.07 (.045)	0.22 (.054)	-0.20 (.059)
Anti-social Punishment	-0.03 (.062)	-0.03 (.053)	0.05 (.049)	-0.05 (.041)		0.43 (.034)	-0.20 (.061)	0.22 (.059)
Pro-social Punishment	0.06 (.058)	0.08 (.050)	0.03 (.048)	0.07 (.045)	0.43 (.033)		-0.04 (.061)	0.04 (.061)
Dislike Having More	0.16 (.053)	0.22 (.047)	0.25 (.057)	0.22 (.054)	-0.20 (.061)	-0.04 (.061)		0.26 (.059)
Dislike Having Less	-0.07 (.054)	-0.05 (.054)	-0.04 (.061)	-0.20 (.059)	0.22 (.059)	0.04 (.061)	0.26 (.059)	

Notes: Bootstrapped standard errors from 10,000 simulations in parentheses. Colors in heatmap change with each 0.05 of magnitude of correlation.

of inequality aversion. These clusters are characterized by high within-cluster correlations and low correlations between measures in different clusters. To make these clusters visually apparent, we present the correlation matrix in the form of a “heat map” where the shade of red indicates the magnitude of the correlation.<sup>27</sup>

The first correlation in the first cluster—0.86 between the two reciprocity measures—is a useful example for interpretation. There are distinctions between those who are more and less reciprocal when a partner is more or less generous, resulting in a less than perfect correlation. Yet, the predominant behavioral distinction, reflected in the very high correlation, is how reciprocal someone is in all conditions. This overarching behavior is (empirically) related to both trust and altruism, although more closely to the former. We note that although some readers may have anticipated some of these correlations, the fact that the literature distinguishes between, say, different forms of reciprocity, indicates that these anticipations are not universally shared.

The second cluster contains the two punishment measures.<sup>28</sup> Like other clusters, they are highly correlated with each other, but with few other measures. However, these measures are characterized by both an extensive and intensive margin. The extensive margin is whether or not someone punishes, and the intensive how much punishment a person metes out. This leads to obvious questions about how these two margins contribute to the overall relationship. The correlations on both margins are roughly equal. About 2/3 of participants engage in pro-social punishment, whereas only 1/3 engage in anti-social punishment. However, almost everyone who engages in anti-social punishment also engages in pro-social punishment. Of those that engage in both types of punishment, there is a correlation of  $\sim 0.4$  in the amount they choose to spend on punishment of both parties. Both the extensive and intensive margin are poorly related to other measures of social preferences.

The third cluster contains the distributional preferences measures. It features the weakest intra-cluster correlation. Moreover, both measures in this cluster—Dislike Having More and Dislike Having Less—are moderately correlated with other econographics. These mod-

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<sup>27</sup>The shading is driven by a concave function of magnitude so that there is more differentiation between magnitudes of 0 and 0.25 than there is between 0.25 and 1. We also show both the upper and lower part of the symmetric correlation matrix.

<sup>28</sup>As a reminder, both refer to costly punishments meted out on participants in a sender-receiver game where the sender sent his or her entire endowment, and the receiver returned none. Pro-social punishment refers to the amount used to punish the receiver, and anti-social punishment refers to the amount used to punish the sender.



Table 4: Principal Components Analysis of Social Preferences

	<i>Generosity</i>	<i>Punishment</i>	<i>Inequality Aversion</i>	Unexplained
Reciprocity: Low	<b>0.52</b>	0.04	-0.04	0.27
Reciprocity: High	<b>0.52</b>	0.03	0.00	0.27
Altruism	<b>0.40</b>	0.02	0.11	0.54
Trust	<b>0.49</b>	-0.02	-0.07	0.35
Anti-social Punishment	-0.03	<b>0.70</b>	0.05	0.24
Pro-social Punishment	0.07	<b>0.64</b>	-0.03	0.37
Dislike Having More	0.17	-0.24	<b>0.65</b>	0.28
Dislike Having Less	-0.13	0.17	<b>0.75</b>	0.22
Percent of Variation	34%	19%	16%	32%

Notes: First three principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

erate correlations extend to some measures in the risk domain, leading to this component combining with risk preferences in the analysis of all 21 econographics in Section 4.4.

Before turning to the principal components analysis (PCA), it is worth noting that there is very little in the literature that indicates which specific correlations we should, and should not, find. This could lead to a lot of plausible storytelling. For example, one might believe that the dislike of having more (one form of inequality aversion) entirely motivates altruism, and a dislike of having less (another form) motivates punishment. Yet, these patterns are not present in our data. Alternatively, one might expect that the predominant feature of distributional preferences is the presence or absence of a preference for equality, which would generate the observed correlation between Dislike of Having More and Dislike of Having Less. Whatever one’s priors (or lack thereof), these results should be informative. We discuss what we believe are the biggest takeaways for theory in Section 6.

The PCA of these correlations shows the same patterns: results appear in Table 4. Three components are suggested for inclusion under parallel analysis—see Appendix Figure C.1 for the scree plot. Together, these three components explain 68% of the variation in the eight measures of social preferences we explore here.

These three components have fairly obvious interpretations. The first is generosity in behaviors that directly influence the well-being of another person. The second is a general

affinity for punishment. The third and final component captures both types of inequality aversion. The components are thus named accordingly: *Generosity*, *Punishment*, and *Inequality Aversion*.

Overall, these results show there is a clear structure to the connections between different social preferences. These measures tend to group in clusters with high within-cluster correlations and low across-cluster correlation, leading to the latent structure shown by the PCA.

### 4.3 Links Between Risk Attitudes

In this subsection, we show that models of decision-making under risk and uncertainty may be too parsimonious. As shown by the clear clusters in Table 5, ambiguity aversion and compound lottery aversion group together, blurring the line between risk and uncertainty. Moreover, we find two clusters of risk attitudes: one related to WTA and the other to WTP.

The easiest cluster to interpret contains Ambiguity and Compound Lottery Aversion.<sup>29</sup> This high correlation is consistent with extant empirical work (Halevy, 2007; Dean and Ortoleva, 2019; Gillen et al., 2019) and theoretical observation (Segal, 1990; Dean and Ortoleva, 2019). Note that both of these measures are essentially unrelated to measures of risk attitudes towards ordinary lotteries, suggesting either a delineation between risk and uncertainty, with compound lotteries grouped with ambiguity aversion, or the lack of a clean line between risk and ambiguity.

The two remaining clusters all contain risk attitudes, and suggest two separate aspects of risk preferences. The first contains Risk Aversion: Gain, Loss, Gain/Loss, and WTA; and the second includes Risk Aversion: CR Certainty, CR Lottery, and WTP. Crucially, there are low correlations between the measures in different clusters (with the exception of the relationship between Risk Aversion: Losses and WTP—correlated 0.3, meaning that risk aversion using WTP is negatively correlated with risk aversion over losses). This suggests that there are two separate aspects of risk aversion that are largely unrelated to each other, or to preferences under uncertainty. Note that these two different aspects are both within a fairly narrow domain: the valuation of risky lotteries. Thus, unlike psychology research, which has shown

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<sup>29</sup>Note that these come from the certainty equivalent of a draw from an ambiguous (Ellsberg) or compound urn minus the certainty equivalent from a 50/50 risky urn. In order to prevent spurious correlation, we adapt the ORIV procedure as described in Section 3.2.

Table 5: ORIV Correlations of Risk Measures

	WTA	Risk Aversion: Gains	Risk Aversion: Losses	Risk Aversion: Gain/Loss	WTP	Risk Aversion: CR Certain	Risk Aversion: CR Lottery	Ambiguity Aversion	Compound Lottery Aversion	Endowment Effect	Common Ratio
WTA		-0.66 (.051)	-0.27 (.076)	-0.58 (.054)	-0.11 (.069)	-0.03 (.069)	0.12 (.071)	0.14 (.064)	0.09 (.067)	-	-0.14 (.070)
Risk Aversion: Gains	-0.66 (.051)		0.39 (.070)	0.60 (.049)	0.04 (.071)	0.09 (.068)	-0.13 (.070)	0.03 (.057)	0.01 (.061)	-0.49 (.055)	0.20 (.055)
Risk Aversion: Losses	-0.27 (.077)	0.39 (.070)		0.65 (.056)	0.30 (.076)	-0.19 (.074)	-0.15 (.077)	0.06 (.065)	0.01 (.065)	-0.38 (.060)	-0.04 (.064)
Risk Aversion: Gain/Loss	-0.58 (.054)	0.60 (.049)	0.65 (.055)		0.19 (.069)	-0.14 (.070)	-0.21 (.072)	-0.02 (.058)	-0.02 (.072)	-0.53 (.047)	0.06 (.069)
WTP	-0.11 (.069)	0.04 (.070)	0.30 (.076)	0.19 (.069)		-0.45 (.041)	-0.28 (.061)	0.01 (.054)	0.00 (.059)	-	-0.17 (.060)
Risk Aversion: CR Certain	-0.03 (.070)	0.09 (.069)	-0.19 (.074)	-0.14 (.069)	-0.45 (.041)		0.41 (.060)	-0.11 (.063)	-0.05 (.068)	0.26 (.053)	-
Risk Aversion: CR Lottery	0.12 (.071)	-0.13 (.071)	-0.15 (.072)	-0.21 (.072)	-0.28 (.060)	0.41 (.060)		0.01 (.059)	0.06 (.069)	0.26 (.055)	-
Ambiguity Aversion	0.14 (.063)	0.03 (.057)	0.06 (.065)	-0.02 (.058)	0.01 (.055)	-0.11 (.063)	0.01 (.058)		0.47 (.080)	0.09 (.055)	-0.11 (.073)
Compound Aversion	0.09 (.068)	0.01 (.060)	0.01 (.064)	-0.02 (.072)	0.00 (.058)	-0.05 (.068)	0.06 (.068)	0.47 (.066)		0.06 (.057)	-0.10 (.083)

Notes: Bootstrapped standard errors from 10,000 simulations in parentheses. Colors in heatmap with each 0.05 of magnitude of correlation.

Table 6: Principal Components Analysis of Risk Preferences

	<i>Risk Aversion: WTA</i>	<i>Risk Aversion: WTP</i>	<i>Uncertainty</i>	Unexplained
WTA	<b>-0.51</b>	-0.09	0.12	0.32
Risk Aversion: Gains	<b>0.55</b>	0.13	0.05	0.25
Risk Aversion: Losses	<b>0.39</b>	-0.21	0.09	0.45
Risk Aversion: Gain/Loss	<b>0.53</b>	-0.09	0.02	0.22
WTP	0.03	<b>-0.54</b>	-0.02	0.43
Risk Aversion: CR Certain	0.08	<b>0.62</b>	-0.06	0.29
Risk Aversion: CR Lottery	-0.05	<b>0.48</b>	0.11	0.53
Ambiguity Aversion	-0.01	-0.04	<b>0.69</b>	0.27
Compound Lottery Aversion	0.01	0.04	<b>0.69</b>	0.28
Percent of Variation	29%	21%	17%	34%

Notes: First three principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

differences in risk preferences across outcome domains, these patterns suggest differences in preferences within the domain of money lotteries—the most studied domain in choice theory, and experimental and behavioral economics.

These two separate aspects of risk preference align with WTA and WTP, respectively, giving this division substantive and theoretical importance. We find that certainty equivalents fall into one cluster while lottery equivalents into another. This is in line with prior research finding that different fixed elements in an MPL lead to different degrees of risk aversion. Here, as in the literature, this is consistent with the fixed element of the MPL acting as a reference point (Sprenger, 2015). Our results add, first, that these groups are largely uncorrelated with each other, and, second, that they naturally align with WTA and WTP. WTA—where the lottery is explicitly the reference point—naturally aligns with certainty equivalents: measures where the lottery is fixed. Similarly, WTP aligns with measures with a fixed monetary amount. At the same time, as we discuss in Section 6, and more in depth in Chapman et al. (2021), our data are difficult to reconcile with current theoretical explanations. The most visible signature of this difficulty in Table 5 is the lack of an obvious component of loss aversion positively related to WTA and negatively related to Risk Aversion: Gain/Loss.

Before proceeding, it is worth discussing the two measures in the correlation table that we do not include in the clusters or in the following principal components analysis: the endowment effect and the common ratio effect. Both measures are the difference of two other measures we already include: for example, the endowment effect is the difference of the WTA and WTP. Thus, including all three measures in a PCA makes little sense. As described above, our own recent research compiles extensive evidence that WTA and WTP should be treated as separate behavioral measures (Chapman et al., 2021). Thus, we include the component measures (WTA and WTP) rather than the endowment effect. We take the same approach with the common ratio measures: entering the constituent measures into the correlation and principal components analysis.<sup>30</sup>

Once again, the principal components analysis in Table 6 confirms the visual patterns in the correlation table. The first three components explain 66% of the variation in the nine measures of risk preferences considered here. The first and the second clearly capture different aspects of risk attitudes. Following the discussion above, we use the names *Risk Aversion: WTA* and *Risk Aversion: WTP* to describe them. The third encompasses more complex lotteries, which may induce uncertainty: hence we use the title *Uncertainty*.

#### 4.4 Putting it all Together

Analyzing all 21 econographics together leads to six components. The structure of components from the risk and social domains are preserved, however, one social and one risk component combine. The sixth component is comprised of the three overconfidence measures. Time preferences load on several components; most heavily on the Punishment component.

Analyzing risk and social preferences, together with overconfidence and time preferences, is straightforward because there are few important relationships between risk and social preferences. However, they are not completely unrelated. One of the social preference components (Inequality Aversion) combines with a risk-preference component (Risk Aversion:WTP). Otherwise, the structure in the previous two subsections is largely unaltered.

The first and second components in Table 7 are similar to the first components in the social and risk analyses, respectively. Thus, we retain the same names. Patience loads moderately on both components: more generous and less risk averse people are more patient.

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<sup>30</sup>Entering the common ratio measure directly in the principal component analysis produces qualitatively similar results to Table 6; see Table E.2.

Table 7: Principal Components Analysis of All Measures

	<i>Generosity</i>	<i>Risk Aversion: WTA</i>	<i>Inequality Aversion: WTP</i>	<i>Overconfidence</i>	<i>Impulsivity</i>	<i>Uncertainty</i>	Unexplained
Reciprocity: Low	<b>0.50</b>	0.00	0.00	0.02	0.05	-0.01	0.29
Reciprocity: High	<b>0.51</b>	0.01	-0.02	-0.07	0.06	0.04	0.27
Altruism	<b>0.39</b>	0.06	0.03	0.04	0.05	-0.03	0.54
Trust	<b>0.47</b>	0.01	-0.06	0.12	-0.04	-0.05	0.33
Anti-social Punishment	-0.01	0.00	0.03	0.05	<b>0.67</b>	-0.03	0.26
Pro-social Punishment	0.09	-0.05	-0.04	-0.08	<b>0.60</b>	0.03	0.42
Dislike Having More	0.22	0.00	<b>0.26</b>	-0.18	-0.11	0.17	0.57
Dislike Having Less	-0.06	-0.10	<b>0.40</b>	-0.02	0.22	0.07	0.48
WTA	-0.04	<b>-0.50</b>	-0.06	0.06	0.04	0.10	0.31
Risk Aversion: Gains	0.04	<b>0.55</b>	0.11	0.03	-0.05	0.06	0.23
Risk Aversion: Losses	-0.09	<b>0.36</b>	-0.16	0.05	0.16	0.09	0.42
Risk Aversion: Gain/Loss	-0.01	<b>0.50</b>	-0.11	-0.04	-0.01	0.03	0.24
WTP	-0.01	0.01	<b>-0.47</b>	0.14	0.01	-0.02	0.45
Risk Aversion: CR Certain	-0.02	0.10	<b>0.53</b>	0.07	-0.02	-0.11	0.38
Risk Aversion: CR Lottery	-0.03	-0.03	<b>0.42</b>	0.11	-0.03	0.04	0.56
Ambiguity Aversion	0.04	0.00	-0.03	-0.03	0.05	<b>0.70</b>	0.25
Compound Lottery Aversion	-0.07	0.01	0.02	0.08	-0.07	<b>0.65</b>	0.30
Overestimation	0.04	-0.03	0.01	<b>0.50</b>	0.10	0.01	0.54
Overplacement	0.03	0.04	0.16	<b>0.49</b>	-0.09	-0.05	0.56
Overprecision	-0.01	-0.02	-0.07	<b>0.62</b>	-0.02	0.06	0.33
Patience	0.19	-0.19	-0.05	-0.05	<b>-0.28</b>	0.05	0.65
Percent of Variation	14%	13%	11%	8%	8%	7%	40%

Notes: First six principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

The third component is of particular interest. It combines the measures included in the second risk component, Risk Aversion: WTP, with those of the third social component, Inequality Aversion. This suggests that some aspects of social and risk preferences are distinct, while others are more related. Moreover, this relationship follows a clear pattern building upon the previously identified components: Inequality Aversion combines with Risk Aversion: WTP. We note that both of these components deal with aversions to spreads—in possible payoffs or distributional assignments. Thus, we conjecture that a similar form of caution may lead participants to both dislike entering conditions of risk, as well as generically dislike unequal allocations. Note, however, that this pertains only to entering a situation of risk (WTP for a lottery), rather than leaving one (WTA). In line with its constituent parts, we call this component *Inequality Aversion/WTP*.

The fourth component loads heavily on all three overconfidence measures; naturally, we call it *Overconfidence*. Given existing work on the conceptual distinctions between different types of overconfidence, the fact that they would load on a single component may not have been ex ante obvious (Moore and Healy, 2008).

The fifth is the social component Punishment, to which (a bit of) time preferences have been added. People who score highly on this component enjoy both pro- and anti-social punishment, and are impatient. Thus, we dub this new component, *Impulsivity*.

The sixth component is essentially the same as the third risk component. Thus, we continue to call this component *Uncertainty*.

There are a number of ways to examine the robustness of our results. We discuss two here—changing the number of components, and using other methods to extract latent variables—and another—adding and removing variables—in Section 8.

The Inequality Aversion / WTA component is robust to adding or subtracting components. Reducing the number of components to five, as shown in Online Appendix Table E.3, causes the Uncertainty component to combine with both the Overconfidence and Impulsivity components, but otherwise leaves the structure of the components qualitatively unchanged. Perhaps more interesting is what happens when we add components. With both seven or eight components—the latter recommended by the Kaiser Criterion—the Overconfidence component splits into two, as shown in Online Appendix Tables E.4 and E.5. Both of these overconfidence components contain overprecision, but one contains overestimation, and the other, overplacement. Adding an eighth component further splits the Generosity component

into two: with the reciprocity measures on one component, and the altruism and trust measures on the other. The rest of the components are largely unchanged by the inclusion of additional components.

Using other ways of computing the correlation matrix, or other latent dimension recovery techniques, has very little effect on our results, as shown in Appendix Tables E.6 through E.9. In particular, our results are robust to using unweighted measures or the average of measures (rather than ORIV) to compute the correlation matrix. Moreover, one can use Spearman rank-order correlations, which are robust to non-linear relationships and outliers, to compute the correlation matrix, and this also results in very little change to the structure shown in Table 7. Finally, one can use factor analysis (on the average of measures), rather than principal components analysis. This, too, produces qualitatively similar results.

Overall the underlying structure of our data can be summarized with four points. First, six interpretable and separate components explain 60% of the variance of our twenty-one variables. Second, the social and risk components do not change from the earlier analyses in Tables 4 and 6, but two components—Risk Aversion: WTP and Inequality Aversion—collapse into one. Third, time preferences are captured by several of the existing components, possibly signaling that they are related to many elements, although none particularly strongly. Fourth, overconfidence measures are largely separate, and captured by a single component.

## 5 Econographics and Other Measures

In this section, we examine the correlation between our econographic components, cognitive abilities, and demographics. This has two potential benefits. First, it may give us clues as to the the psychological processes underlying these clusters. Second, it may also provide useful information about economic preferences. For example, understanding how the components relate to demographics might be useful in understanding how preferences might change as a population becomes older or better educated.

### 5.1 Econographics and Cognitive Measures

There are several important relationships between our six econographic components and the two measures of cognition on our survey: IQ and the three-question Cognitive Reflection Test



(CRT; Frederick, 2005). These two cognitive ability measures have an ORIV correlation of  $> 0.8$ .<sup>31</sup>

Our IQ measure consists of two classes of questions drawn from the International Cognitive Ability Resource, a public domain intelligence measure (ICAR; Condon and Revelle, 2014). We chose three matrix reasoning questions, similar to Raven’s Progressive Matrices. In these questions, participants determined which of a set of possibilities correctly completed a graphic pattern. We also constructed a second battery of three questions based on 3-D rotations: a drawing of a cube was shown, and participants had to identify which of a set of six other drawings of a cube were compatible. These questions were chosen to be of progressively greater difficulty in order to try to capture a variation in fluid intelligence in the general population.<sup>32</sup>

There are several important relationships between our six econographic components and IQ and the Cognitive Reflection Test (CRT), shown in Table 8. Once again, we present the correlations in the form of a heatmap where shade is indicative of magnitude. Unlike prior sections that were primarily concerned with magnitudes, here, positive and negative correlations have very different interpretations. As such, we distinguish between signs of the correlation using different colors: blue for negative, and green for positive. Consistent with the high correlation between the IQ and CRT measures, the general pattern of correlations is largely the same: higher cognitive ability is positively correlated with Generosity, and negatively correlated with Inequality Aversion/WTP, Overconfidence, and Impulsivity.<sup>33</sup>

These results show that, in general, the underlying behavioral components that we have

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<sup>31</sup>Note that our survey also contained two instantiations of the “beauty contest” measure of strategic sophistication (Nagel, 1995). Unfortunately, as in Snowberg and Yariv (2021), we found that the distribution of responses is relatively uniform, with a spike in responses at 50 (out of 100), as shown in Figure E.1. It is difficult to discern from this pattern who is strategically sophisticated. Thus, we did not include these measures in our analysis.

<sup>32</sup>Condon and Revelle (2014) contains the percent of people that typically get each of their questions correct. We used this information to attempt to cut the population at the 20th, 40th, 60th, 80th, 90th, and 95th percentile. We were largely successful, with 21% answering no questions correctly, 24% answering one correctly, 24% answering two correctly, 15% answering three correctly, 7% answering four correctly, 5% answering five correctly, and 4% answering all six correctly.

<sup>33</sup>Note that perceptions of how well one did on the IQ tasks are used to calculate one of the three main components of the Overconfidence measure, and also have some weight in the other components, which might result in some spurious correlations. This is unlikely to be a matter of great concern as the pattern of correlations with the CRT is largely the same. Moreover, this is consistent with the so-called Kruger and Dunning effect 1999. That is, poor performers also have poor meta-cognition and think their performance is not as relatively bad as it is, leading to negative correlation between confidence and performance (see also, Krueger and Mueller, 2002).

Table 8: Correlations with Measures of Cognitive Ability

	IQ	CRT
Generosity	0.13 (.038)	0.09 (.038)
Risk Aversion: WTA	0.02 (.040)	0.04 (.039)
Inequality Aversion/WTP	-0.13 (.043)	-0.11 (.039)
Overconfidence	-0.23 (.050)	-0.12 (.049)
Impulsivity	-0.24 (.033)	-0.25 (.032)
Uncertainty	0.05 (.034)	0.02 (.037)

Notes: Bootstrapped standard errors from 10,000 simulations in parentheses. Colors in heatmap change with each 0.05 of magnitude of correlation. Blue hues indicate negative correlations, green hues positive correlations.

identified are not simple proxies for intelligence. Moreover, there is a more nuanced relationship between our measures of intelligence and the extent to which a participant is “behavioral”—that is, behaves differently than a selfish, risk-neutral expected utility maximizer. While it is the case that lower measured intelligence is related to higher impulsivity and inequality aversion, it is also related to lower generosity. Moreover, risk and uncertainty attitudes are largely unrelated to intelligence, with much of the positive relationship between Inequality Aversion/WTP, Overconfidence, and intelligence being driven by distributional preferences.<sup>34</sup>

## 5.2 Econographics and Demographics

There are interesting relationships in our data between our components and standard economic variables—such as income and education—or demographics—such as age and gender. Table 9 shows these correlations in Panel A, and the same correlations controlling for IQ in Panel B.<sup>35</sup>

<sup>34</sup>Appendix Table E.1 shows the relationships between individual econographics and cognitive measures.

<sup>35</sup>Appendix Table E.1 shows the relationships between individual econographics and demographics.

Table 9: Correlations with Demographics

	Income	Education	Age	Male	Attend Church
Panel A: Correlations					
Generosity	0.10 (.046)	0.06 (.049)	0.09 (.047)	0.04 (.046)	0.15 (.046)
Risk Aversion: WTA	0.00 (.049)	0.06 (.046)	0.03 (.050)	-0.03 (.047)	0.02 (.054)
Inequality Aversion/WTP	0.05 (.044)	-0.03 (.046)	0.20 (.045)	-0.08 (.046)	-0.06 (.054)
Overconfidence	0.00 (.052)	0.01 (.050)	0.04 (.047)	0.10 (.048)	0.12 (.046)
Impulsivity	-0.19 (.045)	-0.22 (.039)	0.05 (.045)	0.04 (.047)	0.05 (.046)
Uncertainty	0.03 (.045)	0.01 (.044)	-0.01 (.043)	-0.04 (.045)	-0.04 (.041)
Panel B: Controlling for IQ					
Generosity	0.08 (.045)	0.04 (.048)	0.11 (.046)	0.04 (.049)	0.16 (.048)
Risk Aversion: WTA	0.00 (.047)	0.06 (.045)	0.04 (.049)	-0.03 (.047)	0.02 (.053)
Inequality Aversion/WTP	0.07 (.044)	-0.01 (.046)	0.18 (.044)	-0.07 (.046)	-0.06 (.049)
Overconfidence	0.04 (.052)	0.07 (.049)	0.01 (.044)	0.12 (.048)	0.12 (.049)
Impulsivity	-0.16 (.047)	-0.18 (.041)	0.01 (.045)	0.06 (.047)	0.04 (.044)
Uncertainty	0.02 (.045)	0.00 (.043)	0.00 (.043)	-0.05 (.046)	-0.04 (.040)

Notes: Bootstrapped standard errors from 10,000 simulations in parentheses. Colors in heatmap change with each 0.05 of magnitude of correlation. Blue hues indicate negative correlations, green hues positive correlations.

Greater education and income are associated with higher Generosity and lower Impulsivity, as was the case with higher cognitive ability. This may be unsurprising, as education and income are often associated with higher cognitive ability. However, as Panel B of Table 9 shows, the relationships with education and income remain even after controlling for IQ. There is no association between income or education and Overconfidence or Inequal-

ity Aversion/WTP; however, once we control for cognitive abilities, we obtain that higher education is (slightly) correlated with higher Overconfidence. That is, even though more education increases participants' own ability and knowledge, it seems to increase their perceptions of these skills even more. This suggests, if anything, that education is not a "cure" for non-standard preferences.<sup>36</sup>

Older people are more generous and men are more overconfident, in accordance with the literature.<sup>37</sup> These results only become stronger once we control for cognitive abilities. Additionally, older people exhibit a higher Inequality Aversion/WTP score, while men are slightly lower. As with cognitive abilities, we find that neither demographic is related to the other risk component, Risk Aversion: WTA, adding nuance on the relation between these aspects and risk preferences. Lastly, frequency of church attendance is positively associated with Generosity. It is also positively associated with Overconfidence, although not statistically significantly at conventional levels.

## 6 Relation to Theory

As suggested in prior sections, commonly used theories can account for some, but not all, of the patterns in our data, even within the social or risk domain. In this section, we spell out the correlations that would be predicted by different theories, and the points of difference between theories and our empirical patterns.

### 6.1 Social Preferences

Theories of social preferences can capture some of the moments in our data, but no single theory can explain all the patterns we observe. Here we consider four common theories of social preferences: A simple model of altruistic preferences, inequality aversion (Fehr and Schmidt, 1999), reciprocity (Rabin, 1993), and a model incorporating the latter two features plus social welfare (Charness and Rabin, 2002).<sup>38</sup>

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<sup>36</sup>Further, we note that Stango and Zinman forthcoming find larger variation in the degree to which they classify people as "behavioral" within educational strata than across them.

<sup>37</sup>Although not significant after Bonferroni correction, it is significant without. As this fact has been previously shown, for example in Ortoleva and Snowberg (2015b), there is little risk of this being a spurious correlation.

<sup>38</sup>Charness and Rabin (2002) is similar to our approach as they parameterize different components of social preference to see which ones have the most predictive power. But they explicitly, "[H]ave not tested

A simple model of altruistic preferences—in which utility is increasing in one’s own monetary payoff and that of others—predicts a positive correlation between our measures of Altruism, Trust, Reciprocity: High, and Reciprocity: Low, in line with the Generosity component (see Tables 3, 4, and 7). This simple model also predicts that more altruistic people would have to be paid more to move to an uneven split in which the other person got less, but paid less to move to a split in which the other person got more. That is, it also predicts that measures in the Generosity component should be positively correlated with Dislike Having More and negatively related to Dislike Having Less. There is some evidence for such a pattern in our data, however, it does not reflect the most parsimonious reading, as expressed by PCA. Moreover, a model of altruistic preferences would also predict a *negative* correlation between Dislike Having More and Dislike Having Less—the opposite of what we find in our data. The altruistic preferences model also predicts that people would accept a reduction in allocation to move away from an even split to one in which another person got more. In our data, the majority of non-indifferent participants require an *increase* in income (79% of non-indifferent participants in question 1 and 74% in question 2) to move away from the even split, the opposite of the theoretical prediction. Finally, a purely altruistic person would never engage in costly punishment.

Social welfare preferences modify the altruistic preferences model by assuming that the utility a decision-maker gets from the consumption of others grows faster when others are poorer than the decision-maker (Charness and Rabin, 2002). Thus, altruistic behavior is governed by two parameters, one that dictates behavior when a person has more than others, and one when they have less. Like the altruism model, this model also predicts a positive correlation between Altruism, Trust, Reciprocity: High, and Dislike Having More, as these are all situations in which the decision-maker will (choose to) have more than the other person. In contrast, Reciprocity: Low would be governed by a separate parameter, which would also determine Dislike Having Less and lead to a negative correlation between these two measures. This is not the primary pattern we see in the data. As with altruistic preferences, a person of this type would never engage in costly punishment.

Perhaps the most common characterization of social preferences is Fehr and Schmidt’s model of inequality aversion. This adds two parameters to the standard model: one codes

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for individual differences and correlation across games, and neither our analysis nor our model deals with heterogeneity of subject preferences” (p. 849). See also Bruhin et al. (2019).

how much the decision-maker dislikes having more than other people, and the other how much they dislike having less. As with social welfare preferences, the first of these parameters governs behavior in Altruism, Trust, Reciprocity: High, and Dislike Having More, while Dislike Having Less is governed by the second parameter. The model of Bolton and Ockenfels (2000) would imply similar correlations. Again, this does not match the most parsimonious reading of our data. Moreover, a model of inequality aversion would predict that no money should be returned in the Reciprocity: Low measure, but that is not what happens empirically. Inequality-averse participants might engage in Pro-social Punishment (as this would reduce inequality) but not Anti-social Punishment (as this would increase inequality). This is inconsistent with the strong positive correlation we find between these two measures. The single punishment factor is also inconsistent with evidence from cross-cultural studies on public goods games with punishment, such as Gächter et al. (2010), in which cultures tend to bifurcate between pro-social and anti-social punishment.

Other important theories of social preferences are not tested in our survey. Intention-based models, such as Rabin (1993) only apply to one of our choice environments—the trust game—that has a second player whose actions affect both players’ payoffs. Ideal tests of these models need to collect actions, beliefs about others’ actions, and second-order beliefs (for example, Charness and Dufwenberg, 2006), which are hard to measure, especially given the constraints of our survey (Schotter and Trevino, 2014). While we do not have ideal measures of intention, to test reciprocity, a simple comparison of our items can plausibly lead to a negative correlation between Reciprocity: High (in which participants have been treated fairly) and Reciprocity: Low (in which they have been treated unfairly). This is in contrast to what we find in our data.<sup>39</sup>

## 6.2 Risk Preferences

Our results exhibit inconsistencies with standard models of preferences under risk and uncertainty. These models typically separate risk and uncertainty, with the distinction being whether probabilities of outcomes are known or unknown (Arrow, 1951; Gilboa and Marinacci, 2016). The curvature of the utility function, and possibly aspects of non-Expected

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<sup>39</sup>A related class of theories models social image concerns, in which behavior depends delicately on what people think others know about their behavior (for example, Lazear et al., 2012; Andreoni and Bernheim, 2009). Establishing common knowledge and credibility are crucial for the validity of such measures, but very difficult to do in representative surveys.

Utility, should affect behavior in both domains, while beliefs and ambiguity aversion should only impact choice under uncertainty. In our data, the classic delineation between risk and uncertainty fails to hold: attitudes towards compound lotteries, where probabilities are known, are related to ambiguity aversion, where they are not. This empirical relationship has been shown before, with different possible explanations (Halevy, 2007; Dean and Ortoleva, 2019; Gillen et al., 2019). Segal (1990) proposes that ambiguous prospects are treated like compound lotteries, in which case probability weighting could explain both phenomena. This theory would further imply a link between Ambiguity/Compound Lottery Aversion and the Common Ratio effect. However, we do not observe this relationship in our data. An alternative is to argue that compound lotteries are complex objects that people perceive as ambiguous (Dean and Ortoleva, 2019). In either case, these results point to a different structure, where compound lotteries are associated with prospects with unknown probabilities.

The split into two separate types of risk preferences for objective, non-compound, risk is also difficult to reconcile with models that treat risk aversion as a unitary phenomenon—for example, driven by the curvature of a utility function. As discussed at length in Chapman et al. (2021), the specific clusters we find can be interpreted using the framework of Sprenger (2015), in which question structure serves as an implicit frame and determines reference points. In particular, Sprenger shows that certainty and probability equivalents yield different levels of risk aversion. He explains this finding by appealing to the different reference points these frames induce. In line with this hypothesis, the two clusters in our data are related to the two explicit frames we administer in the WTA and WTP tasks: certainly equivalents are correlated with WTA, while probability equivalents are correlated with WTP.

A natural question is whether this pattern of correlations is in line with existing theories of reference dependence. Cumulative Prospect Theory (CPT) and the Kőszegi Rabin (KR) model predict specific relationships between WTA, WTP, and the Endowment Effect, and the three measures Risk-Aversion: Gains, Losses, and Gain/Loss. While deriving these specific predictions and testing them is beyond the scope of this paper, Chapman et al. (2021) shows that the predicted relationships fail to hold, and in some cases are in the opposite direction of predictions from reference-dependence models. Thus, while the reference point induced in these questions is of clear first-order importance, existing theories of reference dependence

cannot account for the overall pattern of correlations observed in our data.

## 7 Literature

Before reviewing the few papers that, like ours, examine the correlations between large sets of behavioral regularities, we note that there is a small but significant literature focusing on the connections between only two or three of them. These papers primarily focus on the relationships between different risk preferences, or risk and time preferences; a couple also study the links between social preferences and risk. Broadly speaking, where these studies overlap with our work, these studies find similar relationships. We discuss this literature in greater detail in Appendix D.

Another group of papers studies the relationship between risk and/or time preferences and cognitive ability (Burks et al. 2009; Dohmen et al. 2010; Benjamin et al. 2013—see Dohmen et al. 2018 for a recent review). In aggregate, the literature suggests that higher cognitive ability is associated with more risk aversion. Our results suggest some nuance, as we find different relationships between cognitive ability and the two components of ordinary risk attitudes we identify.

Our work is differentiated from the studies described above by its large representative sample, and measures of multiple behavioral regularities. This allows us to identify the complete correlational structure without worries about confounds from differing study designs or populations. However, two contemporaneous and complementary projects also have multiple behavioral measures and use representative populations. Stango and Zinman (forthcoming) measure a broad set of 17 behavioral factors in a representative sample of U.S. adults to study, in part, the relationships between them. Falk et al. (2018) survey 80,000 adults across 80 countries to document patterns of six behaviors. Two additional studies use multiple elicitations in non-representative populations.<sup>40</sup> These studies differ from ours in both purpose and implementation: they study different measures, do not use incentives, and do not take

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<sup>40</sup>Becker et al. (2012) measure risk aversion, time preferences, trust, and altruism using incentivized questions in a representative sample. However, they do not report the correlations between these behaviors, as their goal is to study the links between these behaviors and personality. Burks et al. (2009) make use of a large scale experiment carried out on a group of newly recruited truck drivers. The authors use parametric methods to measure risk aversion, short-term and long-term discounting (though a beta-delta model) and behavior in a sequential two-person prisoner’s dilemma (similar to our trust game). They find a statistically significant (though quantitatively small) relationship between risk attitudes, patience, and sender behavior in the trust game: more patient people tended to be less risk averse and send more in the trust game.



the steps that we do to eliminate the attenuating effects of measurement error.

Stango and Zinman (forthcoming) is the closest to our study in spirit, and presents a different, but complementary, way of thinking about increasing the parsimony of behavioral economics. They measure a wide-ranging set of behaviors, including statistical biases such as the gambler’s fallacy, but excluding social preferences. We focus on a narrower set of (primarily) preference measures.<sup>41</sup> Additionally, they focus on relationships with consumer behavior to help select dimensions that are important for consumer decision-making. Their broader focus allows their results to be useful to theories of cognition, whereas our narrower focus makes our results more directly related to choice and decision theory.

Falk et al. (2018) collect a subset of the measures we include here: patience, risk, positive reciprocity, punishment, altruism, and trust. Each is measured through a combination of qualitative self-reports and hypothetical-money questions. Three of the the four measures of social preferences—“altruism,” “positive reciprocity,” and “trust”—are highly correlated, a result which is reproduced in our data.<sup>42</sup> The purposes of this paper are simply different than ours. For example, they are interested in a huge range of 80 countries, differences between countries, and cross-country associated variables (such as language and religion). Their measures are also unincentivized, and it is certainly useful to now have two sets of data which vary on this dimension.

An immediate predecessor of this paper is Dean and Ortoleva (2019), which studies relationships between many of the same behaviors in a sample of 180 Brown undergraduates. These students are likely to be less heterogeneous on some dimensions, such as cognitive ability. Dean and Ortoleva focus more on risk and time, and less on social preferences.

Broadly speaking, to the extent that these studies use similar measures, all three tend to find similar relationships, although that overlap is relatively small: ranging from 15 (Falk et al., 2018) to 45 (Dean and Ortoleva, 2019) of the 210 correlations we examine.

As all of these studies are trying to understand a complicated question of first-order importance to behavioral economics—How are many different behaviors related?—it is natural that there would be some overlap. Because this question is both fundamental and complicated, and involves a large number of possible measures, it is not likely that there will be a

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<sup>41</sup>They use exploratory factor analysis for dimension reduction, which is closely related to PCA. Indeed, our results are relatively unchanged by using factor analysis rather than PCA, see Table E.9.

<sup>42</sup>Correlations between these preferences are generally lower than what we observe (all are below 0.33). This may be due to the attenuating effects of measurement error, or the fact that they examine a more diverse population (across 80 countries).

single, definitive, study soon. Instead, we believe that these many studies trace out different features of this complicated enterprise. The ideal cumulative scientific process, in our opinion, is what this area of research is currently working towards: Creating a set of studies, all with distinctive strengths, that can be generally assessed as results are published.

## 8 Conclusion

We elicit 21 econographics from a representative sample of 1,000 U.S. adults in order to create an empirical basis for an underlying structure of more comprehensive theories of behavioral decision-making. We identify six interpretable econographic components that explain a large fraction of the variance in these 21 econographic variables: Generosity, Risk Aversion: WTA, Inequality Aversion/WTP, Overconfidence, Impulsivity, and Uncertainty. These components suggest that more parsimonious representations than current theories of social preferences are possible, but that canonical theories of risk preferences are perhaps too parsimonious. Moreover, they suggest limited, and nuanced, connections between risk and social preferences. By studying the relationship between the components we identify and cognitive measures and demographics, we document several stylized facts that may be useful for theorizing.

A strength of our study is the number of behaviors included in our analysis. However, the behaviors we included were limited both by survey time and by the current literature.<sup>43</sup> These, in turn, present limits for our analyses. A nuanced view of these limitations comes from thinking about what would happen if we had included more, or fewer, measures, which we do in the remainder of this paper. This exercise also speaks to the robustness of our results.

Including an extremely similar elicitation to “Risk Aversion: Gains” has little qualitative effect on our conclusions, as shown in Table E.10. This extremely similar measure loads heavily on the second component in Table 7, and this component now becomes the first (in terms of percent of variation explained).<sup>44</sup> Thus, the ranking of components may respond

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<sup>43</sup>For example, we would have liked to have a measure of motivated reasoning. However, we were unable to find such a measure in the literature, and our design principles, discussed in Section 2, dictated that we not attempt to invent our own.

<sup>44</sup>This measure, Risk Aversion: Urn is described in Footnote 13. Note that this measure also loads on the uncertainty component, although the correlations that lead to this loading are spurious, as Risk Aversion: Urn is used in constructing both of the measures that comprise the uncertainty component.

to the inclusion or exclusion of measures. As such, we have not attached meaning to the ordering in the text.

Removing Dislike Having More reduces the number of components in Table 4 to two, but has little qualitative effect on the overall analysis of Table 7, as shown in Table E.11.<sup>45</sup> Dislike Having Less still combines with the variables in the Risk Aversion: WTP component to form Inequality Aversion/WTP. Thus, it appears that minor perturbations are not particularly consequential, although these exercises have little to say about larger changes.

The robustness here stems from the fact that most correlations between measures—displayed, for example, in Table 3 and 5—are either large or close to zero.<sup>46</sup> There are a few middling correlations—such as those between Dislike Having More and other social preference measures. These are the likely sources of fragility in PCA. Adding a variable with a number of middling correlations may cause parallel analysis to suggest the inclusion of an additional component, and this inclusion may lead to extensive changes in existing components. As such, the components are useful for making sense of the large correlation matrices in our analysis. However, the correlation matrices themselves are the most robust—and consequential—part of the analysis, as these are the fundamental patterns upon which latent variable models, such as PCA, are built.

Overall, this discussion suggests that our main conclusion—that there is an underlying structure to our measures that is informative for theorizing—is robust. However, we note that our findings relate to fairly specific domains of economic behaviors: choices over money lotteries, time, the distribution of resources between two people, and beliefs about oneself and others. Adding more measures might create broader clusters that increase the average number of measures per component, or lead to a more diffuse set of underlying dimensions. Moreover, examining the relationship between clusters and consumer behavior, as in (Stango and Zinman, forthcoming), might, in the context of specific applications, increase the salience of particular relationships and/or the importance of some components. Whatever the outcome of such explorations, it is worth noting that a failure to reduce all economically-relevant behaviors to just a handful of components is not a failure of behavioral economics. Chemistry has been incredibly successful with more than 100 elements. In econographics, we have

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<sup>45</sup>The number of components is determined, as before, using parallel analysis. In analyzing just the social preference measures, the Dislike Having Less loads on the punishment component.

<sup>46</sup>The preceding two paragraphs are not comprehensive tests of robustness. Rather, they illustrate patterns in the data to build intuition about how PCA responds to small changes in which econographics are included.

barely started down that path. It should be our goal to accurately and adequately describe economic behavior with no more, and no fewer, components than necessary. This study takes a step in that direction.

## References

- Abdellaoui, Mohammed, Peter Klibanoff, and Lætitia Placido**, “Experiments on Compound Risk in Relation to Simple Risk and to Ambiguity,” *Management Science*, 2015, *61* (6), 1306–1322.
- Abdi, Hervé and Lynne J Williams**, “Principal Component Analysis,” *Wiley Interdisciplinary Reviews: Computational Statistics*, 2010, *2* (4), 433–459.
- Ahn, David, Syngjoo Choi, Douglas Gale, and Shachar Kariv**, “Estimating Ambiguity Aversion in a Portfolio Choice Experiment,” *Quantitative Economics*, 2014, *5* (2), 195–223.
- Allais, Maurice**, “Le Comportement de l’Homme Rationnel devant le Risque: Critique des Postulats et Axiomes de l’Ecole Americaine,” *Econometrica*, 1953, *21* (4), 503–546.
- Andersen, Steffan, Glenn W. Harrison, Morten I. Lau, and E. Elisabet Rutström**, “Eliciting Risk and Time Preferences,” *Econometrica*, 2008, *76* (3), 583–618.
- Andreoni, James and B Douglas Bernheim**, “Social image and the 50–50 norm: A theoretical and experimental analysis of audience effects,” *Econometrica*, 2009, *77* (5), 1607–1636.
- and **Charles Sprenger**, “Estimating time Preferences from Convex Budgets,” *The American Economic Review*, 2012, *102* (7), 3333–3356.
- and – , “Risk Preferences Are Not Time Preferences,” *American Economic Review*, 2012, *102* (7), 3357–3376.
- Arrow, Kenneth J.**, “Alternative Approaches to the Theory of Choice in Risk-taking Situations,” *Econometrica*, 1951, *19* (4), 404–437.
- Augenblick, Ned, Muriel Niederle, and Charles Sprenger**, “Working Over Time: Dynamic Inconsistency in Real Effort Tasks,” *The Quarterly Journal of Economics*, 2015, *130* (3), 1067–1115.
- Azrieli, Yaron, Christopher P. Chambers, and Paul J. Healy**, “Incentives in experiments: A Theoretical Analysis,” *Journal of Political Economy*, 2018, *126* (4), 1472–1503.
- Beattie, Jane and Graham Loomes**, “The Impact of Incentives upon Risky Choice Experiments,” *Journal of Risk and Uncertainty*, March 1997, *14* (2), 155–168.
- Becker, Anke, Thomas Deckers, Thomas Dohmen, Armin Falk, and Fabian Kosse**, “The Relationship between Economic Preferences and Psychological Personality Measures,” *Annual Review of Economics*, 2012, *4* (1), 453–478.
- Bendor, Jonathan and Piotr Swistak**, “The Evolution of Norms,” *American Journal of Sociology*, 2001, *106* (6), 1493–1545.
- Benjamin, Daniel J. Sebastian A. Brown, and Jesse M. Shapiro**, “Who is Behavioral? Cognitive Ability and Anomalous Preferences,” *Journal of the European Economic Association*, 2013, *11* (6), 1231–1255.
- Berg, Joyce, John Dickhaut, and Kevin McCabe**, “Trust, Reciprocity, and Social History,” *Games and Economic Behavior*, 1995, *10* (1), 122–142.
- Bernheim, B. Douglas**, “The Good, the Bad, and the Ugly: A Unified Approach to Behavioral Welfare Economics,” *Journal of Benefit-Cost Analysis*, 2016, *7* (1), 12–68.
- Bolton, Gary E. and Axel Ockenfels**, “ERC: A Theory of Equity, Reciprocity, and Competition,” *The American Economic Review*, 2000, *90* (1), 166–193.
- Bordalo, Pedro, Nicola Gennaioli, and Andrei Shleifer**, “Salience Theory of Choice

- under Risk,” *The Quarterly Journal of Economics*, 2012, 127 (3), 1243–1285.
- , – , and – , “Salience and Consumer Choice,” *Journal of Political Economy*, 2013, 121 (5), 803–843.
- Bruhin, Adrian, Ernst Fehr, and Daniel Schunk**, “The Many Faces of Human Sociality: Uncovering the Distribution and Stability of Social Preferences,” *Journal of the European Economic Association*, 2019, 17 (4), 1025–1069.
- Burks, Stephen V. Jeffrey P. Carpenter, Lorenz Goette, and Aldo Rustichini**, “Cognitive Skills affect Economic Preferences, Strategic Behavior, and Job Attachment,” *Proceedings of the National Academy of Sciences*, 2009, 106 (19), 7745–7750.
- Butler, David and Graham Loomes**, “Imprecision as an Account of the Preference Reversal Phenomenon,” *The American Economic Review*, 2007, 97 (1), 277–297.
- and – , “Imprecision as an Account of Violations of Independence and Betweenness,” *Journal of Economic Behavior & Organization*, 2011, 80 (3), 511–522.
- Carlsson, Fredrik, Dinky Daruvala, and Olof Johansson-Stenman**, “Are People Inequality-Averse, or Just Risk-Averse?,” *Economica*, 2005, 72 (287), 375–396.
- Cason, Timothy N. and Charles R. Plott**, “Misconceptions and Game Form Recognition: Challenges to Theories of Revealed Preference and Framing,” *Journal of Political Economy*, 2014, 122 (6), 1235–1270.
- Cerreia-Vioglio, Simone, David Dillenberger, and Pietro Ortoleva**, “Cautious Expected Utility and the Certainty Effect,” *Econometrica*, 2015, 83 (2), 693–728.
- Chakraborty, Anujit, Yoram Halevy, and Kota Saito**, “The Relation between Behavior under Risk and Over Time,” *American Economic Review: Insights*, 2020, 2 (1), 1–16.
- Chakravarty, Sujoy and Jaideep Roy**, “Recursive Expected Utility and the Separation of Attitudes towards Risk and Ambiguity: An Experimental Study,” *Theory and Decision*, 2009, 66 (3), 199–228.
- Chang, Beryl and Fabrizio Ghisellini**, *Behavioural Economics: An Unfinished Business*, New York: Palgrave MacMillan, 2018.
- Chapman, Jonathan, Pietro Ortoleva, Erik Snowberg, Mark Dean, and Colin Camerer**, “Willingness-To-Pay and Willingness-To-Accept are Probably Less Correlated than You Think,” 2021. California Institute of Technology, *mimeo*.
- Charness, Gary and Martin Dufwenberg**, “Promises and Partnership,” *Econometrica*, 2006, 74 (6), 1579–1601.
- and **Matthew Rabin**, “Understanding Social Preferences with Simple Tests,” *The Quarterly Journal of Economics*, 2002, 117 (3), 817–869.
- , **Uri Gneezy, and Brianna Halladay**, “Experimental Methods: Pay One or Pay All,” *Journal of Economic Behavior & Organization*, 2016, 131, 141–150.
- Chetty, Raj**, “Behavioral Economics and Public Policy: A Pragmatic Perspective,” *American Economic Review*, 2015, 105 (5), 1–33.
- Cohen, Jonathan, Keith Marzilli Ericson, David Laibson, and John Myles White**, “Measuring time preferences,” *Journal of Economic Literature*, 2020, 58 (2), 299–347.
- Cohen, Michèle, Jean-Marc Tallon, and Jean-Christophe Vergnaud**, “An Experimental Investigation of Imprecision Attitude and its Relation with Risk Attitude and Impatience,” *Theory and Decision*, July 2011, 71 (1), 81–109.
- , **Jean-Yves Jaffray, and Tanios Said**, “Experimental Comparison of Individual Behav-

- ior under Risk and under Uncertainty for Gains and for Losses,” *Organizational Behavior and Human Decision Processes*, 1987, 39 (1), 1–22.
- Condon, David M. and William Revelle**, “The International Cognitive Ability Resource: Development and Initial Validation of a Public-domain Measure,” *Intelligence*, 2014, 43 (2), 52–64.
- Conlisk, John**, “Why Bounded Rationality?,” *Journal of Economic Literature*, 1996, 34 (2), 669–700.
- Cubitt, Robin, Chris Starmer, and Robert Sugden**, “On the Validity of the Random Lottery Incentive System,” *Experimental Economics*, September 1998, 1 (2), 115–131.
- Dean, Mark and Pietro Ortoleva**, “The Empirical Relationship between Nonstandard Economic Behaviors,” *Proceedings of the National Academy of Sciences of the United States of America*, 2019, 116 (3), 16262–16267.
- DellaVigna, Stefano**, “Psychology and Economics: Evidence from the Field,” *Journal of Economic Literature*, 2009, 47 (2), 315–372.
- Dohmen, Thomas, Armin Falk, David Huffman, and Uwe Sunde**, “Are Risk Aversion and Impatience Related to Cognitive Ability?,” *American Economic Review*, 2010, 100 (3), 1238–1260.
- , – , – , and – , “On the Relationship between Cognitive Ability and Risk Preference,” *Journal of Economic Perspectives*, 2018, 32 (2), 115–34.
- Dunn, Olive Jean**, “Estimation of the Means of Dependent Variables,” *The Annals of Mathematical Statistics*, 1958, 29 (4), 1095–1111.
- , “Multiple Comparisons among Means,” *Journal of the American Statistical Association*, 1961, 56 (293), 52–64.
- Eckel, Catherine C. and Rick K. Wilson**, “Is Trust a Risky Decision?,” *Journal of Economic Behavior & Organization*, 2004, 55 (4), 447–465.
- Ellsberg, Daniel**, “Risk, Ambiguity, and the Savage Axioms,” *The Quarterly Journal of Economics*, 1961, 75 (4), 643–669.
- Enke, Benjamin and Thomas Graeber**, “Cognitive Uncertainty,” 2019. mimeo, Harvard University.
- Epper, Thomas, Helga Fehr-Duda, and Adrian Bruhin**, “Viewing the Future through a Warped Lens: Why Uncertainty Generates Hyperbolic Discounting,” *Journal of Risk and Uncertainty*, December 2011, 43 (3), 169–203.
- Epstein, Ziv, Alexander Peysakhovich, and David G. Rand**, “The Good, the Bad, and the Unflinchingly Selfish: Cooperative Decision-making can be Predicted with High Accuracy when using only Three Behavioral Types,” in “Proceedings of the 2016 ACM Conference on Economics and Computation” ACM 2016, pp. 547–559.
- Ericson, Keith M Marzilli, John Myles White, David Laibson, and Jonathan D Cohen**, “Money Earlier or Later? Simple Heuristics explain Intertemporal Choices better than Delay Discounting Does,” *Psychological Science*, 2015, 26 (6), 826–833.
- Falk, Armin, Anke Becker, Thomas Dohmen, Benjamin Enke, David Huffman, and Uwe Sunde**, “Global Evidence on Economic Preferences,” *The Quarterly Journal of Economics*, 2018, 4 (133), 1645–1692.
- , – , – , **David Huffman, and Uwe Sunde**, “An Experimentally Validated Preference Survey Module,” 2013. University of Bonn, *mimeo*.
- Farhi, Emmanuel and Iván Werning**, “Monetary Policy, Bounded Rationality, and

- Incomplete Markets,” *American Economic Review*, 2019, 109 (11), 3887–3928.
- Fehr, Ernst and Klaus M. Schmidt**, “A Theory Of Fairness, Competition, And Cooperation,” *The Quarterly Journal of Economics*, August 1999, 114 (3), 817–868.
- and **Urs Fischbacher**, “Third-party Punishment and Social Norms,” *Evolution and Human Behavior*, 2004, 25 (2), 63–87.
- Forsythe, Robert, Joel L. Horowitz, N. Eugene Savin, and Martin Sefton**, “Fairness in Simple Bargaining Experiments,” *Games and Economic behavior*, 1994, 6 (3), 347–369.
- Frederick, Shane**, “Cognitive Reflection and Decision Making,” *Journal of Economic Perspectives*, 2005, 19 (4), 25–42.
- Freeman, David, Yoram Halevy, and Terri Kneeland**, “Eliciting Risk Preferences using Choice Lists,” *Quantitative Economics*, 2019, 10 (1), 217–237.
- Friedman, Daniel, R Mark Isaac, Duncan James, and Shyam Sunder**, *Risky Curves: On the Empirical Failure of Expected Utility*, Routledge, 2014.
- Frydman, Cary and Lawrence J Jin**, “Efficient Coding and Risky Choice,” *The Quarterly Journal of Economics*, 2022, 137 (1), 161–213.
- Fudenberg, Drew**, “Advancing Beyond Advances in Behavioral Economics,” *Journal of Economic Literature*, Aug 2006, 44 (3), 694–711.
- Furr, R. Michael**, *Psychometrics: An Introduction*, Sage Publications, 2017.
- Gächter, Simon, Benedikt Herrmann, and Christian Thöni**, “Culture and Cooperation,” *Philosophical Transactions of the Royal Society B: Biological Sciences*, 2010, 365 (1553), 2651–2661.
- García-Schmidt, Mariana and Michael Woodford**, “Are Low Interest Rates Deflationary? A Paradox of Perfect-Foresight Analysis,” *American Economic Review*, 2019, 109 (1), 86–120.
- Gilboa, Itzhak and Massimo Marinacci**, “Ambiguity and the Bayesian Paradigm,” in Horacio Arló-Costa, Vincent F. Hendricks, and Johan van Benthem, eds., *Advances in Economics and Econometrics: Theory and Applications*, Vol. 1 of *Springer Graduate Texts in Philosophy*, Springer, 2016, pp. 385–439.
- Gillen, Ben, Erik Snowberg, and Leeat Yariv**, “Experimenting with Measurement Error: Techniques and Applications from the Caltech Cohort Study,” *Journal of Political Economy*, 2019, 127 (4), 1826–1863.
- Halevy, Yoram**, “Ellsberg Revisited: An Experimental Study,” *Econometrica*, 2007, 75 (2), 503–536.
- Herrmann, Benedikt, Christian Thöni, and Simon Gächter**, “Antisocial Punishment across Societies,” *Science*, 2008, 319 (5868), 1362–1367.
- Hey, John and Jinkwon Lee**, “Do Subjects Separate (or Are They Sophisticated)?,” *Experimental Economics*, September 2005, 8 (3), 233–265.
- Holt, Charles A. and Susan K. Laury**, “Risk Aversion and Incentive Effects,” *American Economic Review*, 2002, 92 (5), 1644–1655.
- Imai, Taisuke, Tom A Rutter, and Colin Camerer**, “Meta-Analysis of Present-Bias Estimation Using Convex Time Budgets,” *The Economic Journal*, 2021, 131 (636), 1788–1814.
- Kahneman, Daniel**, “Maps of Bounded Rationality: Psychology for Behavioral Economics,” *American Economic Review*, 2003, 93 (5), 1449–1475.



- **and Amos Tversky**, “Prospect Theory: An Analysis of Decision under Risk,” *Econometrica*, March 1979, 47 (2), 263–292.
- **, Jack L. Knetsch, and Richard H. Thaler**, “Experimental Tests of the Endowment Effect and the Coase Theorem,” *Journal of Political Economy*, 1990, 98 (6), 1325–1348.
- Karni, Edi and Zvi Safra**, “‘Preference Reversal’ and the Observability of Preferences by Experimental Methods,” *Econometrica*, 1987, 55 (3), 675–685.
- Kerschbamer, Rudolf**, “The Geometry of Distributional Preferences and a Non-parametric Identification Approach: The Equality Equivalence Test,” *European Economic Review*, 2015, 76, 85–103.
- Khaw, Mel Win, Ziang Li, and Michael Woodford**, “Cognitive Imprecision and Small-stakes Risk Aversion,” *The Review of Economic Studies*, 2021, 88 (4), 1979–2013.
- Kőszegi, Botond**, “Behavioral Contract Theory,” *Journal of Economic Literature*, 2014, 52 (4), 1075–1118.
- **and Matthew Rabin**, “Reference-dependent Risk Attitudes,” *American Economic Review*, 2007, 97 (4), 1047–1073.
- Krueger, Joachim and Ross A. Mueller**, “Unskilled, Unaware, or Both? The Better-than-average Heuristic and Statistical Regression Predict Errors in Estimates of Own Performance,” *Journal of Personality and Social Psychology*, 2002, 82 (2), 180–88.
- Kruger, Justin and David Dunning**, “Unskilled and Unaware of it: How Difficulties in Recognizing One’s own Incompetence Lead to Inflated Self-Assessments,” *Journal of Personality and Social Psychology*, 1999, 77 (6), 1121–1134.
- Kurata, Hiroshi, Hiroshi Izawa, and Makoto Okamura**, “Non-Expected Utility Maximizers Behave as if Expected Utility Maximizers: An Experimental Test,” *Journal of Socio-Economics*, 2009, 38 (4), 622–629.
- Lauriola, Marco and Irwin P. Levin**, “Relating Individual Differences in Attitude toward Ambiguity to Risky Choices,” *Journal of Behavioral Decision Making*, 2001, 14 (2), 107–122.
- Lazear, Edward P., Ulrike Malmendier, and Roberto A. Weber**, “Sorting in Experiments with Application to Social Preferences,” *American Economic Journal: Applied Economics*, January 2012, 4 (1), 136–63.
- Levine, David K.**, *Is Behavioral Economics Doomed?: The Ordinary versus the Extraordinary*, Cambridge, UK: Open Book Publishers, 2012.
- Masatlioglu, Yusufcan and Efe A. Ok**, “A Canonical Choice Model with Initial Endowment,” *Review of Economic Studies*, 2014, 81 (2), 851–883.
- Moore, Don A. and Amelia S. Dev**, “Overconfidence,” in Virgil Zeigler-Hill and Shackelford Todd K., eds., *Encyclopedia of Personality and Individual Differences*, Springer, 2020, pp. 3382–3386.
- Moore, Don A. and Paul J. Healy**, “The Trouble with Overconfidence,” *Psychological Review*, 2008, 115 (2), 502–517.
- Nagel, Rosemarie**, “Unraveling in Guessing Games: An Experimental Study,” *The American Economic Review*, 1995, 85 (5), 1313–1326.
- Niederle, Muriel and Lise Vesterlund**, “Do Women Shy Away from Competition? Do Men Compete too Much?,” *Quarterly Journal of Economics*, August 2007, 122 (3), 1067–1101.
- Ortoleva, Pietro and Erik Snowberg**, “Are Conservatives Overconfident?,” *European*

- Journal of Political Economy*, 2015, 40 (B), 333–344.
- and –, “Overconfidence in Political Behavior,” *American Economic Review*, 2015, 105 (2), 504–535.
- Pedroni, Andreas, Renato Frey, Adrian Bruhin, Gilles Dutilh, Ralph Hertwig, and Jörg Rieskamp**, “The Risk Elicitation Puzzle,” *Nature Human Behaviour*, 2017, 1 (11), 803–809.
- Pew Research Center**, *Evaluating Online Nonprobability Surveys*, www.pewresearch.org, May 2016.
- Peysakhovich, Alexander, Martin A. Nowak, and David G. Rand**, “Humans Display a ‘Cooperative Phenotype’ that is Domain General and Temporally Stable,” *Nature Communications*, 2014, 5, 4939.
- Plott, Charles R. and K. Zeiler**, “The Willingness to Pay–Willingness to Accept Gap, The ‘Endowment Effect,’ Subject Misconceptions, and Experimental Procedures for Eliciting Valuations,” *American Economic Review*, 2005, 95 (3), 530–545.
- Quiggin, John**, “A Theory of Anticipated Utility,” *Journal of Economic Behavior & Organization*, 1982, 3 (4), 323–343.
- Rabin, Matthew**, “Incorporating Fairness into Game Theory and Economics,” *The American Economic Review*, 1993, 83 (5), 1281–1302.
- Schechter, Laura**, “Traditional Trust Measurement and the Risk Confound: An Experiment in Rural Paraguay,” *Journal of Economic Behavior & Organization*, 2007, 62 (2), 272–292.
- Schotter, Andrew and Isabel Trevino**, “Belief Elicitation in the Laboratory,” *Annual Review of Economics*, 2014, 6 (1), 103–128.
- Segal, Uzi**, “The Ellsberg Paradox and Risk Aversion: An Anticipated Utility Approach,” *International Economic Review*, 1987, 28 (1), 175–202.
- , “Two-Stage Lotteries without the Reduction Axiom,” *Econometrica*, 1990, 58 (2), 349–377.
- Snowberg, Erik and Leeat Yariv**, “Testing the Waters: Behavior across Subject Pools,” *American Economic Review*, 2021, 111 (2), 687–719.
- Sprenger, Charles**, “An Endowment Effect for Risk: Experimental Tests of Stochastic Reference Points,” *Journal of Political Economy*, 2015, 123 (6), 1456–1499.
- Stango, Victor and Jonathan Zinman**, “We are all Behavioral, More or Less: A Taxonomy of Consumer Decision Making,” *Review of Economic Studies*, forthcoming.
- Stevens, Luminita**, “Coarse Pricing Policies,” *The Review of Economic Studies*, 2020, 87 (1), 420–453.
- Thaler, Richard H.**, “Behavioral Economics: Past, Present, and Future,” *American Economic Review*, 2016, 106 (7), 1577–1600.
- von Gaudecker, Hans-Martin, Arthur Van Soest, and Erik Wengström**, “Heterogeneity in Risky Choice Behavior in a Broad Population,” *The American Economic Review*, 2011, 101 (2), 664–694.
- Weber, Elke U. and Eric J. Johnson**, “Decisions under Uncertainty: Psychological, Economic, and Neuroeconomic Explanations of Risk Preference,” in Paul W. Glimcher, Ernst Fehr, Colin Camerer, and Russell Alan Poldrack, eds., *Neuroeconomics: Decision Making and the Brain*, New York, NY: Elsevier, 2008, pp. 127–144.
- Woodford, Michael**, “Modeling Imprecision in Perception, Valuation, and Choice,” *Annual*

*Review of Economics*, 2020, 12, 579–601.

# Online Appendix—Not Intended for Publication

## A Survey Implementation

### A.1 Measures

This section builds on Section 2 with a more direct description of how we instantiate the various measures discussed there. In the interest of brevity, measures are summarized as succinctly as possible. The specific question wordings, randomization of question ordering, screenshots, and so on can be found in the supplementary materials uploaded with this submission and online at [eriksnowberg.com/wep.html](http://eriksnowberg.com/wep.html). As mentioned in Section 3.2, in order to reduce issues related to measurement error, we elicit many quantities twice, and apply ORIV. We sometimes refer to these second elicitations as “duplicates.”

#### A.1.1 Altruism, Trust, Reciprocity, and Punishment

**Altruism:** Participants are given 6,000 points to allocate, in 1,000 point increments, between themselves and another randomly chosen participant (Forsythe et al., 1994). The amount given to the other participant is our measure of altruism.

**Trust:** Participants play a standard sender-receiver trust game, in both the sender and receiver roles (Berg et al., 1995). As the sender, they are given another stock of 6,000 points to keep or to send to a randomly chosen participant—the receiver—in 1,000 point increments.<sup>1</sup> The amount sent is doubled, and then the receiver can send some portion of the total back, which is again doubled. We treat the amount sent as our measure of trust.

**Reciprocity, High and Low:** We elicit the responses of each participant, acting as a receiver, to each possible quantity of points he or she may receive in the trust game. These are in 2,000 point increments, from 2,000 to 12,000 points, reflecting the six possible amounts that may be sent. As we do not observe the response of the receiver to the lowest possible action (sending 0 points), it is possible that receivers encode sending 1,000 points as a neutral, or even positive, action. As such, we call the response to receiving 2,000 points

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<sup>1</sup>Participants are informed that this randomly chosen participant is different from the one in the dictator game.

“Reciprocity: Low”, with the amount sent back when receiving 4,000 points as the duplicate measure for ORIV. Symmetrically, we label the amount sent back when receiving 12,000 points as “Reciprocity: High,” although it could safely be labeled positive reciprocity. The amount sent back when receiving 10,000 points is the duplicate measure for ORIV.

**Pro- and Anti-social Punishment:** Participants observe another sender-receiver trust game that they do not participate in, in which the sender sent 6,000 points and the receiver sent back 0 points, resulting in payoff of 12,000 for the receiver and 0 for the sender. Participants are given a stock of 4,000 points, which they can use to deplete the points of the sender and the receiver. Each point spent depletes the points of the targeted individual (the sender or the receiver) by six points. They can spend up to 2,000 points punishing each participant, and any unused points are kept. Participants are told (truthfully) that no one will have the opportunity to punish their own behavior, and that if their responses are randomly selected for implementation, no one else will have the opportunity to punish the sender or receiver.<sup>2</sup>

The amount used punishing the receiver is our measure of pro-social punishment, as this can be seen as a participant attempting to enforce the pro-social norm of trust. We label the amount used to punish the sender anti-social punishment (as extensively documented in Herrmann et al., 2008).<sup>3</sup>

### A.1.2 Distributional Preferences

Our design for eliciting distributional preferences follows Kerschbamer (2015), which nests several popular models including Fehr and Schmidt (1999) and Charness and Rabin (2002). This technique results in two measures. In both of these, the participant decides on an allocation of points between themselves and another, randomly-matched, participant.

**Dislike Having Less:** In each line of the MPL, the participant decides between an even allocation (say, 4,000 points each), or an uneven allocation that favors another participant. In the uneven allocation, the other participant always receives the same amount, which is greater than both the even allocation and the allocation to the participant making the

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<sup>2</sup>This measure was made incentive compatible by applying the chosen punishments to additional participants recruited to take part in a lab experiment.

<sup>3</sup>Note that in the context of public goods games, pro-social punishment is sometimes referred to in the literature as “third-party costly punishment.” (Bendor and Swistak, 2001; Fehr and Fischbacher, 2004).

decision (say, 6,500 points). Each line of the MPL increments the amount to the participant making the choice (from, say, 1,600 points to 6,400 points). This defines how much the decision-making participant is willing to forgo in order to avoid having less than the other participant. Note that this quantity can be negative, in the sense that being willing to take, say, 3,000 points in order to ensure the other person will get 6,500 points means they are willing to forgo some points to ensure that the other person has more. Note, however, that a negative quantity here does not imply that the participant dislikes having more.

**Dislike Having More:** In this case, the participant is deciding between an even split, and another split that would give the other participant a lower total (say, 1,500 points) and the decision-making participant more (say, between 1,600 and 6,500 points). Once again, this quantity can be negative if the decision-making participant is willing to diminish their own points in order to ensure that they get more.

### A.1.3 Risk and Loss Aversion

This and the following three sets of measures require participants to price lotteries in various ways. All lotteries involve only two possible payoffs, and most assign 50% probability to each. The specific point values of the lotteries used to measure each quantity can be found in Table A.1. Question wordings and screenshots can be found in the supplementary materials uploaded with this submission and at [eriksnowberg.com/wep.html](http://eriksnowberg.com/wep.html).

**Risk Aversion: Gains:** Participants are asked to give their certainty equivalent, using an MPL, for two lotteries containing either gains, or a gain and a zero amount.

**Risk Aversion: Losses:** Certainty equivalents are elicited for two lotteries that contain either only losses, or a loss and a zero amount.

**Risk Aversion: Gain/Loss:** Certainty equivalents are elicited for two lotteries that contain a loss and a gain of equal magnitude.

### A.1.4 Common Ratio

The common ratio effect suggests non-linear preferences over, or misperceptions of, probabilities (Allais, 1953).

**Risk Aversion: CR Certain:** The participant gives the lottery equivalent for a fixed certain amount  $a$ . That is, for a fixed probability  $\alpha$ , this question identifies  $b$  such that  $a \sim \alpha b + (1 - \alpha)0$  for that participant. ( $b$  is sometimes referred to as a *gain-equivalent*.)

**Risk Aversion: CR Lottery:** Fixing a ratio  $x > 1$ , we use the same  $a$  and  $\alpha$  as above, and the participant gives lottery equivalent  $c$  such that  $(1/x)a + (1 - 1/x)0 \sim (\alpha/x)c + (1 - \alpha/x)0$ .

**Common Ratio:** Under Expected Utility,  $b = c$ . We measure the common ratio effect as the difference between the two measures,  $b - c$ .

### A.1.5 Uncertainty

**Risk Aversion: Urn:** This is similar to “Risk Aversion: Gains” however, the participant is told that his or her payment is tied to a random draw from a jar with 100 balls, divided 50/50 between two colors, rather than being told there is a 50% probability of a certain payment. One color is tied to a high payoff, and the other to a zero payoff. In this and the following two measures the participant is asked to choose which color they would like to result in the high payoff. Empirically, this measure is highly correlated with “Risk Aversion: Gains”, and is primarily used to difference out the risk component from ambiguous and compound urns.<sup>4</sup>

**Compound Lottery Aversion:** The payoffs from this lottery are also tied to a draw from an urn. However, participants are told that the exact composition of the urn is not yet determined. Rather, a number between 0 and 100 will be drawn, and this will be the number of balls of a specific color in the jar, and the other color will make up the remaining balls. The certainty equivalent for this lottery, minus the certainty equivalent for Risk Aversion: Urn is the measure of compound lottery aversion (Halevy, 2007).

**Ambiguity Aversion:** Here, the payoffs are tied to a draw from an Ellsberg (1961) urn—that is, an urn with an unknown proportion of balls of each color. The certainty equivalent for this lottery, minus the certainty equivalent for Risk Aversion: Urn is the measure of ambiguity aversion.

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<sup>4</sup>It is also used to illustrate an aspect of principal components analysis in Section 8.

Note that both Compound Lottery Aversion and Ambiguity Aversion difference out the same quantity. If this quantity is measured with error, this can create spurious correlation between Compound Lottery Aversion and Ambiguity Aversion. This issue, and our solution, are discussed in Section 3.2.

#### **A.1.6 Endowment Effect**

**WTP:** Participants are given a stock of points they can divide between keeping and buying a lottery ticket, using an MPL. The amount they are willing to spend on the ticket is used as their willingness to pay (WTP).

**WTA:** Participants are endowed with a lottery ticket, and their selling price is elicited using an MPL. The amount they are willing to sell the ticket for is their willingness to accept (WTA).

**Endowment Effect:** The endowment effect is measured as the difference between WTA and WTP for the same lottery ticket (Kahneman et al., 1990; Plott and Zeiler, 2005).

The same two lottery tickets are used in both WTA and WTP. In order to avoid consistency effects, one of WTA or WTP is elicited near the beginning of the survey, and the other is elicited near the end. Whether WTA or WTP comes first for a given participant is randomly determined.

#### **A.1.7 Overconfidence**

Overconfidence can be divided into three types: overestimation, overplacement, and overprecision (Moore and Healy, 2008). Overestimation refers to a person's estimate of her performance on a task (versus her actual performance). Overplacement refers to her perceived performance relative to other participants (versus her real relative performance). Overprecision refers to a belief that one's information is more precise than it actually is.

**Overestimation:** After each block of three questions on the in-survey IQ test (described in Section 5.1) the participant was asked how many questions they thought they got correct (out of three). This, minus the actual number correct is used as our measure of overestimation.



Table A.1: Lotteries Used for Risk Measures

	Lottery 1	Lottery 2
Panel A: Fixed Lotteries		
Risk Aversion: Gains	$0.5 * 0 \oplus 0.5 * 5,000$	$0.5 * 1,000 \oplus 0.5 * 4,000$
Risk Aversion: Losses	$0.5 * -5,000 \oplus 0.5 * 0$	$0.5 * -4,000 \oplus 0.5 * -1,000$
Risk Aversion: Gain/Loss	$0.5 * -5,000 \oplus 0.5 * 5,000$	$0.5 * -4,000 \oplus 0.5 * 4,000$
Risk Aversion: Urn	$0.5 * 0 \oplus 0.5 * 10,000$	$0.5 * 0 \oplus 0.5 * 8,000$
Compound Lottery Aversion	$0.5 * 0 \oplus 0.5 * 10,000$	$0.5 * 0 \oplus 0.5 * 8,000$
Ambiguity Aversion	Ambiguous $\{0, 10,000\}$	Ambiguous $\{0, 8,000\}$
WTA	$0.5 * 1,000 \oplus 0.5 * 10,000$	$0.5 * 2,000 \oplus 0.5 * 8,000$
WTP	$0.5 * 1,000 \oplus 0.5 * 10,000$	$0.5 * 2,000 \oplus 0.5 * 8,000$
Panel B: Variable Lotteries		
Common Ratio: High Probability		
Fixed (a)	2,500	4,000
Variable	$0.2 * 0 \oplus 0.8 * b$ $b \in [2,200, 7,000]$	$0.25 * 0 \oplus 0.75 * b'$ $y \in [3,600, 10,000]$
Common Ratio: Low Probability		
Fixed (a)	$0.75 * 0 \oplus 0.25 * 2,500$	$0.8 * 0 \oplus 0.2 * 4,000$
Variable	$0.8 * 0 \oplus 0.2 * c$ $c \in [2,200, 7,000]$	$0.85 * 0 \oplus 0.15 * c'$ $c' \in [3,600, 10,000]$

Notes: Possible values for  $b, c \in \{2200, 2500, 2800, 3100, 3400, 3700, 4000, 4300, 4600, 4900, 5200, 5500, 5800, 6100, 6400, 6700, 7000\}$ , and  $b', c' \in \{3600, 4000, 4400, 4800, 5200, 5600, 6000, 6400, 6800, 7200, 7600, 8000, 8400, 8800, 9200, 9600, 10000\}$ .

**Overprecision:** Our measures of overprecision and overplacement follow Ortoleva and Snowberg (2015b,a). Participants are asked two factual questions: about the year the telephone was invented, and the current unemployment rate. After each factual question, the participant is asked for a qualitative assessment of the accuracy of their answer. This residual from a regression of this measure on a fourth-order polynomial of the participant's accuracy gives overprecision.

**Overplacement:** After giving their subjective perception of how accurate their answers were, the participant is asked to give their perception about their accuracy compared to a random group of 100 other survey takers. This answer is the perceived percentile of their

accuracy. The actual percentile of their accuracy is subtracted from this number to give our measure of overplacement.

Note that as both overprecision and overplacement refer to the same factual questions, this will create correlated measurement error between them. This can create spurious correlation between the two measures. This issue, and our solution, is discussed in Section 3.2.

#### **A.1.8 Patience (Time Preferences)**

Patience is measured using the standard “Money Earlier or Later” (MEL) paradigm (for example, Marzilli Ericson et al., 2015). Participants were offered 4,000 points in 45 days, and, in the duplicate, 6,000 points in 90 days. In each row of an MPL, the participant could select the delayed payment, or an alternative amount immediately, where the alternative amount varied between rows.

## **B Machine Learning**

As noted in various points in the text, a few readers have suggested applying machine learning techniques to our data, especially clustering algorithms. As we also noted in the text, these methods failed to produce additional insight. This brief appendix is intended to give readers an understanding of why these methods were not useful in summarizing our data. It is not a comprehensive summary of all of the machine learning techniques we tried, rather, it shows, through the use of clustering algorithms, why those techniques are not useful in this context. There are two causes: there are no strong clusters in our data, and reasons related to standard “curse of dimensionality” problems prevent the identification of more subtle clusters.

In order to provide some intuition, we graph our data, although visualizing a 21-dimensional space is challenging. We approach visualization in two ways. First, we use the first two principal components identified in the text as axes. By construction, these components contain the most information about the underlying data. We also use multi-dimensional scaling, a common visualization technique that tries to preserve the distance between points while projecting them onto a lower dimensional space. The axes produced by this process are not directly interpretable. The figures in this section should provide some insight into why we

did not include any figures of raw data in the text: reducing 21-dimensional data to two dimensions loses a lot of information.

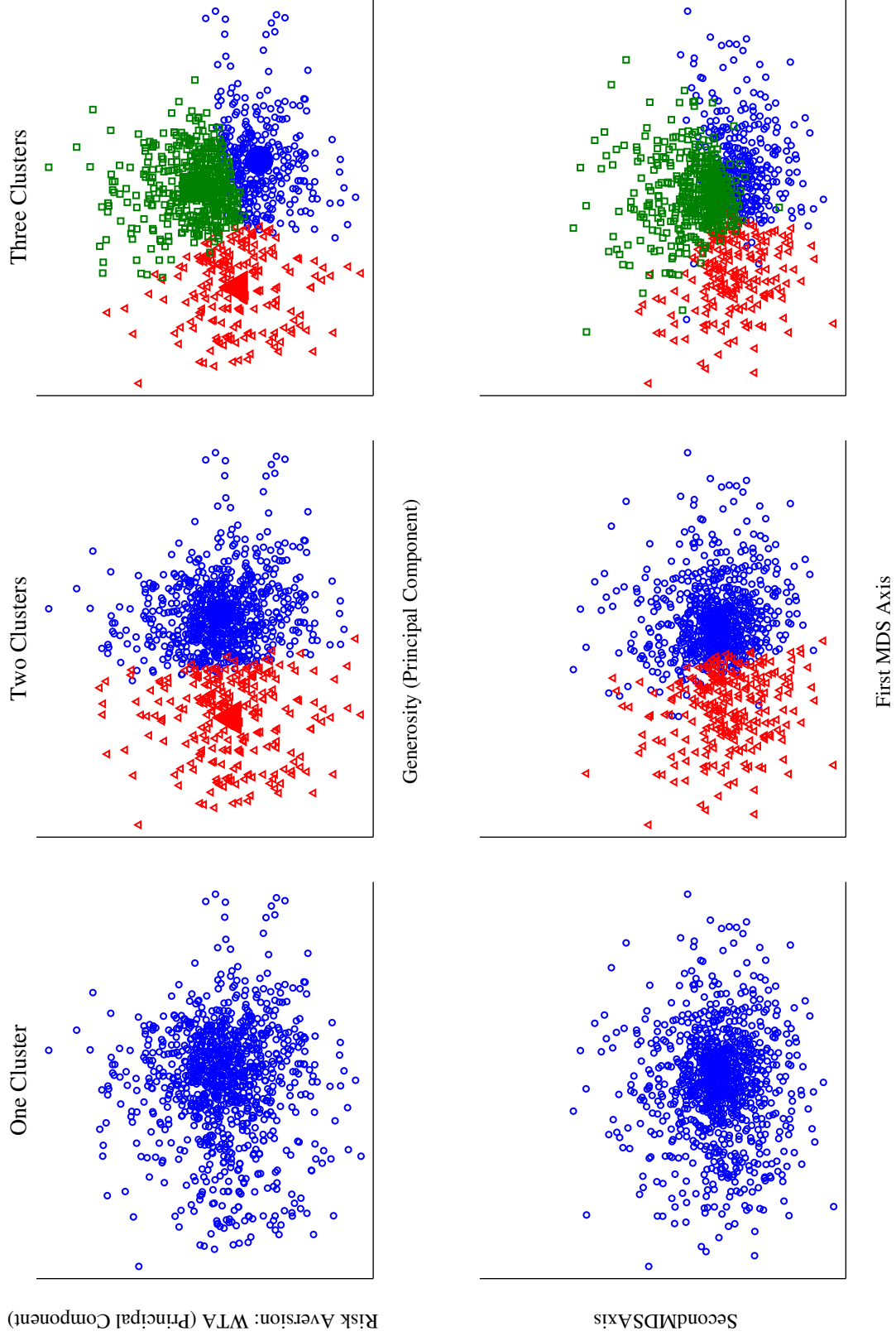
Using standard procedures, such as the so-called “elbow method”, unsupervised clustering algorithms (whether using k-medians or k-means) produce an optimal number of clusters: in our case one cluster, illustrated in the first panel of Figure B.1. This figure plots the data along the first two principal components, averaging measures when we have more than one for a given attribute (such as risk aversion, etc.) Visual inspection should make it obvious why the optimal number of clusters is one: the data look to be distributed according to a single-mode joint distribution, such as a multi-variate normal. We note that the optimal clustering scheme takes into account information from all 21 dimensions, not just the two principal components that are displayed.

Forcing the clustering algorithm to include two or three clusters, as in the second and third panels of Figure B.1, seems to just impose arbitrary dividing lines on the single cluster in the first panel. Using a decision-tree approach to understand the classification shows that the most important determinant of the two clusters is whether or not someone is extremely risk loving on the Risk Aversion: Gain / Loss elicitation. Deeper decision trees provide just as little insight. Substantially similar results are observed by illustrating these three clustering schemes using MDS, as in the bottom three panels of Figure B.1.

Narrowing our focus to a particular domain, such as social preferences, does not help. Even with only our 8 social preference measures, the optimal number of clusters is still one, although the data is “lumpier,” as displayed in Figure B.2. Forcing two clusters does not produce much insight. The clusters are largely defined by whether or not someone falls on one side or the other of the median level of Generosity (the first principal component of the social analysis). This dividing line goes through what appears to be the densest part of the distribution, rather than defining two largely distinct clusters.

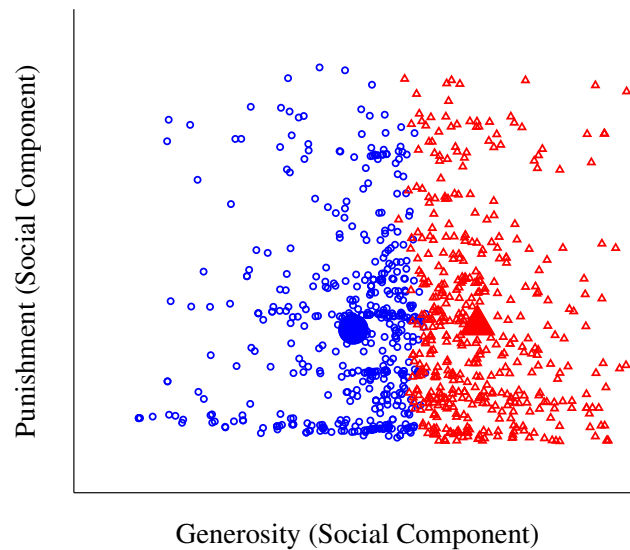
While it is relatively clear that there are no strong clusters in our data, our data would not be optimal for finding more subtle clusters. To understand why, it is instructive to consider the dimensionality of the problem. While 21 dimensions may seem small compared to our 1,000 data points, in the realm of machine learning this is simply not the case. To understand why, suppose we took our 21 continuous measures and divided them along the median value of that measure. This would produce  $2^{21} \approx 2,000,000$  cells. Even with such a coarse classification, the 1,000 datapoints fill 979 cells. Restricting to the domain of social

Figure B.1: The optimal number of clusters in our data is one.



Notes: Top three panels show scatter plots of data in the space defined by the first two principal components of all 21 measures. Color and shape indicates cluster. Large, filled-in shape indicates centroid of a cluster. Bottom three panels show scatter plots of the data in a space defined by multi-dimensional scaling (MDS).

Figure B.2: Restricting focus to social preferences does not improve the performance of clustering algorithms.



Notes: Scatter plots of data in the space defined by the first two of the eight social preference measures. Color and shape indicate cluster. Large, filled-in shape indicates centroid of a cluster.

preferences, with  $2^8 = 256$  cells, the data is still spread out across 133 of them.

It is worth noting that PCA is an often-used tool in machine learning. That is, displaying our main results does take advantage of a machine learning technique, just a particularly simple one that is understood by some economists. The visualizations here build on that analysis, although only in a limited way. However, trying to cluster participants on only their principal component scores (a six-dimensional space) also results in one optimal cluster. This should be unsurprising, as the first six principal components include most of the information in all 21 dimensions.

As noted at the beginning of this appendix, we also tried other types of machine learning analyses, such as outlier detection. Moreover, we also tried building clustering and decision-tree analyses on top of our principal components by, say, taking only one variable from each cluster identified by PCA. These different procedures do not produce additional insight.

Figure C.1: Scree Plot for PCA of Social Measures

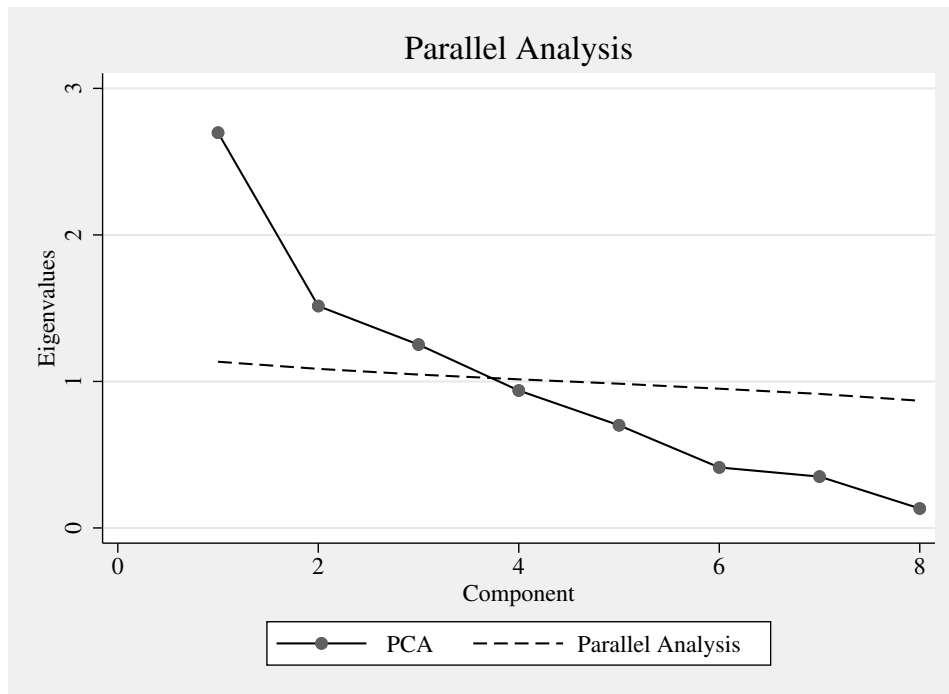


Figure C.2: Scree Plot for PCA of Risk Measures

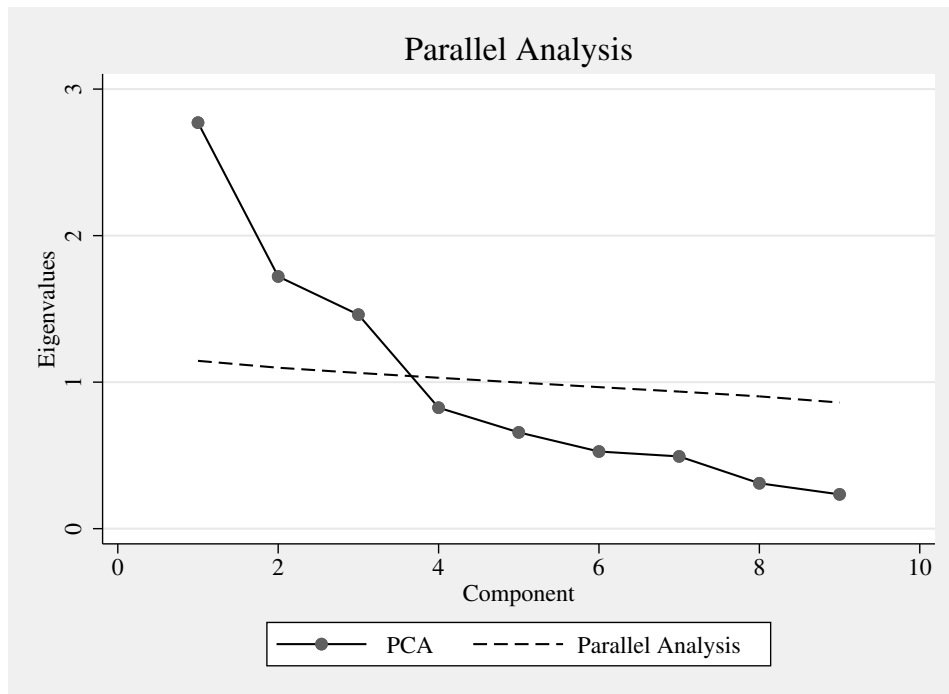
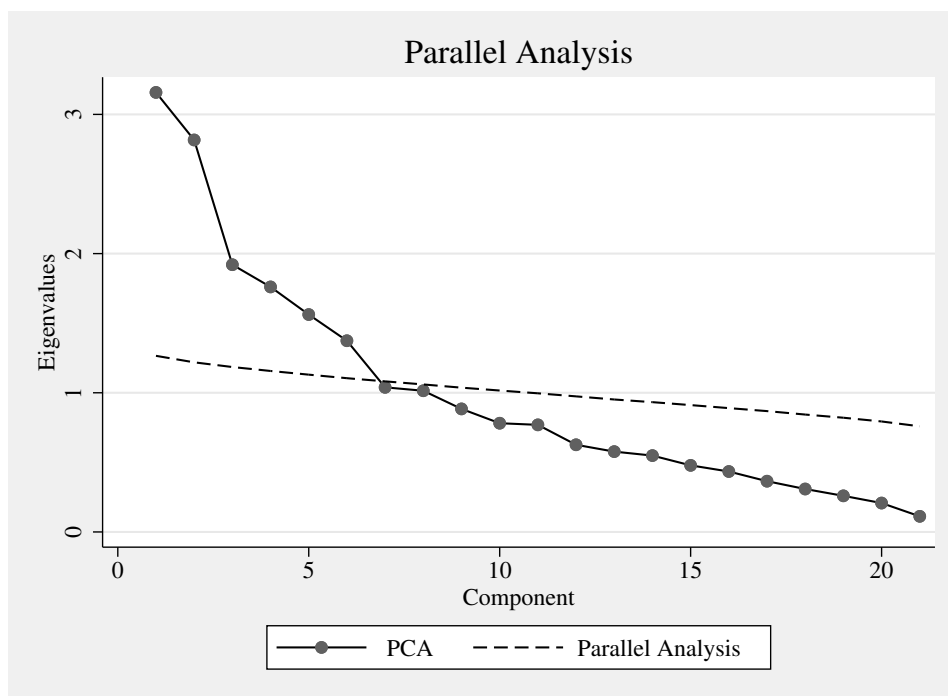


Figure C.3: Scree Plot for PCA of Risk, Social, Overconfidence and Time Preferences



## C Scree Plots of PCAs

## D Additional Literature

There are several laboratory studies of the relationship between different aspects of risk preferences. Many of the findings from these studies extend to our representative population. However, our results create a more nuanced picture of the complicated relationships between different facets of risk attitudes. Pedroni et al. (2017) find that risk attitudes are poorly correlated—and that parameter estimates from a CPT model are unstable—across elicitation tasks based on objective probabilities. This is consistent with our finding that risk attitudes are multi-faceted. A few studies find weak or non-existent relationships between attitudes towards risk and ambiguity/uncertainty (see, for example, Cohen et al., 1987; Lauriola and Levin, 2001; Chakravarty and Roy, 2009; Cohen et al., 2011; Ahn et al., 2014).<sup>5</sup> Our findings are similar, and we also document the relationship between ambiguity and compound lottery

<sup>5</sup>von Gaudecker et al. (2011) document significant heterogeneity in the distribution of preference parameters governing risky choice in a broad population. However, due to the statistical methodology they employ it would be difficult to examine correlations between these parameters.

aversion previously established in a student sample in Halevy (2007).<sup>6</sup>

Several studies describe and test theoretical relationships between risk and time preferences—both of which are related to the curvature of utility functions.<sup>7</sup> Generally, in our data, there is some evidence the two are related, but our results isolate the specific part of risk preferences that are associated with time preferences.

There are fewer studies of the relationships between different aspects of social preferences, or between social preferences and risk. Peysakhovich et al. (2014) report findings similar to ours: cooperative behavior is correlated across a number of settings including the trust and dictator games, but is unrelated to costly punishment. At least one study documents a correlation between risk and inequality aversion, based on the idea that if one hypothetically imagines being endowed with a social position from behind a “veil of ignorance” (à la Rawls) then social inequality creates personal risk (Carlsson et al., 2005). There is conflicting evidence on whether there is a connection between risk attitudes and sender behavior in a trust game (Eckel and Wilson, 2004; Schechter, 2007; Epstein et al., 2016). We find no such relationship.

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<sup>6</sup>Abdellaoui et al. (2015) find a weaker relationship, although Gillen et al. (2019) show that findings of weak relationships in experiments may be due to measurement error. We utilize Gillen et al.’s techniques in order to avoid such issues.

<sup>7</sup>Andersen et al. (2008) find that taking into account risk attitudes reduces estimated discount rates, implying a relationship between risk attitudes and discount rates calculated under the assumption of risk neutrality (as we do in this paper). Andreoni and Sprenger (2012b) find evidence for a failure of *interchangeability*: inter-temporal preferences vary depending on whether the objects of choice are lotteries or certain payments, which they interpret as evidence for different utility functions for risky and risk-free prospects. Finally, Epper et al. (2011) find a significant relationship between decreasing impatience and probability weighting in an experiment that links the two behaviors. Unlike our experiment, they estimate probability weighting parametrically from choices over lotteries, rather than using questions explicitly designed to identify the existence of probability weighting.



## E Other Analyses

Figure E.1: CDF of responses to “Beauty Contest” questions show no signs of strategic sophistication.

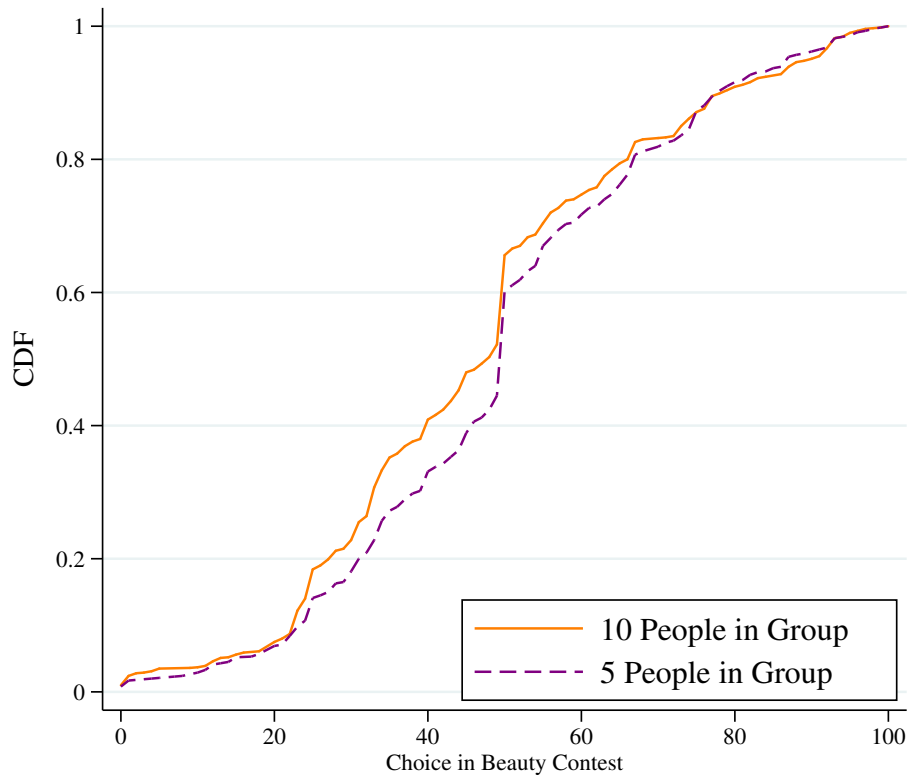


Table E.1: Correlations of Econographics with Cognitive Measures and Demographics

	IQ	CRT	Income	Education	Age	Male	Attend Church
Reciprocity: Low	0.06 (.041)	0.05 (.041)	0.05 (.048)	0.00 (.052)	0.06 (.044)	0.04 (.048)	0.19 (.043)
Reciprocity: High	0.15 (.038)	0.09 (.039)	0.08 (.044)	0.04 (.049)	0.09 (.044)	0.04 (.047)	0.19 (.042)
Altruism	-0.03 (.046)	-0.05 (.047)	0.01 (.047)	0.02 (.046)	0.06 (.047)	0.03 (.046)	0.04 (.046)
Trust	0.13 (.040)	0.07 (.039)	0.06 (.043)	0.07 (.042)	0.05 (.043)	0.06 (.043)	0.08 (.044)
Anti-Social Punishment	-0.23 (.034)	-0.26 (.032)	-0.17 (.047)	-0.21 (.043)	0.03 (.043)	-0.01 (.047)	0.02 (.044)
Pro-Social Punishment	-0.01 (.040)	-0.03 (.041)	-0.07 (.045)	-0.10 (.043)	0.03 (.046)	0.08 (.044)	0.03 (.044)
Dislike Having More	0.17 (.037)	0.14 (.043)	0.17 (.045)	0.09 (.046)	0.10 (.055)	-0.11 (.049)	0.02 (.060)
Dislike Having Less	-0.09 (.041)	-0.08 (.048)	0.03 (.046)	0.00 (.044)	0.13 (.051)	-0.18 (.042)	0.03 (.053)
WTA	-0.04 (.041)	-0.09 (.036)	-0.03 (.052)	-0.10 (.048)	-0.03 (.049)	0.01 (.050)	0.00 (.052)
Risk Aversion: Gains	0.00 (.043)	0.06 (.040)	0.04 (.052)	0.09 (.050)	0.08 (.050)	-0.01 (.050)	-0.01 (.054)
Risk Aversion: Losses	0.00 (.045)	-0.03 (.048)	-0.08 (.056)	-0.03 (.056)	0.01 (.058)	-0.06 (.055)	0.06 (.062)
Risk Aversion: Gain/Loss	0.06 (.043)	0.02 (.044)	0.02 (.050)	0.06 (.051)	-0.01 (.054)	-0.08 (.051)	0.05 (.057)
WTP	0.06 (.046)	0.05 (.046)	-0.08 (.047)	0.04 (.053)	-0.11 (.046)	0.03 (.049)	0.09 (.051)
Risk Aversion: CR Certain	-0.01 (.041)	0.03 (.040)	0.08 (.049)	0.02 (.047)	0.16 (.044)	0.03 (.048)	0.03 (.048)
Risk Aversion: CR Lottery	-0.11 (.047)	-0.11 (.044)	-0.05 (.048)	-0.04 (.048)	0.08 (.051)	-0.04 (.050)	-0.08 (.051)
Ambiguity Aversion	0.05 (.040)	0.01 (.041)	0.02 (.047)	0.00 (.047)	-0.02 (.046)	0.02 (.049)	-0.03 (.045)
Compound Lott. Aversion	0.01 (.040)	0.03 (.043)	0.00 (.050)	0.00 (.053)	-0.01 (.054)	-0.06 (.053)	-0.07 (.046)
Overestimation	-0.62 (.061)	-0.22 (.059)	-0.09 (.063)	-0.11 (.059)	0.20 (.060)	0.12 (.064)	0.18 (.062)
Overplacement	-0.04 (.089)	-0.01 (.067)	0.11 (.066)	0.05 (.075)	0.06 (.075)	-0.07 (.077)	0.10 (.085)
Overprecision	0.03 (.062)	-0.03 (.060)	0.10 (.063)	0.13 (.064)	-0.07 (.061)	0.16 (.061)	0.10 (.069)
Patience	0.23 (.038)	0.21 (.039)	0.16 (.051)	0.16 (.046)	0.04 (.050)	-0.08 (.054)	0.06 (.056)

Table E.2: Principal Components Analysis of Risk Preferences with Common Ratio Measure

	Risk Aversion: WTA	Risk Aversion: WTP	Uncertainty	Unexplained
WTA	<b>-0.50</b>	0.09	0.11	0.33
Risk Aversion: Gains	<b>0.54</b>	0.08	-0.17	0.24
Risk Aversion: Losses	<b>0.37</b>	0.04	0.39	0.35
Risk Aversion: Gain/Loss	<b>0.52</b>	0.01	0.15	0.22
WTP	0.05	-0.09	<b>0.66</b>	0.39
Common Ratio	0.20	-0.09	<b>-0.59</b>	0.47
Ambiguity Aversion	0.00	<b>0.70</b>	0.02	0.27
Compound Lottery Aversion	0.01	<b>0.69</b>	-0.03	0.28
Percent of Variation	33%	19%	17%	32%

Notes: First three principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

Table E.3: Using 5 components, causes Uncertainty to merge with both Overconfidence and Impulsivity (from Table 7)

	<i>Generosity</i>	<i>Risk Aversion:</i> <i>WTA</i>	<i>Inequality Aversion:</i> <i>WTP</i>	<i>Uncertainty/Overconfidence</i>	<i>Uncertainty/Impulsivity</i>	Unexplained
Reciprocity: Low	<b>0.49</b>	-0.01	0.00	0.02	0.03	0.30
Reciprocity: High	<b>0.49</b>	0.00	-0.02	-0.03	-0.01	0.30
Altruism	<b>0.39</b>	0.06	0.03	0.02	0.05	0.55
Trust	<b>0.48</b>	0.02	-0.05	0.07	-0.01	0.34
Anti-Social Punishment	-0.02	-0.02	0.04	0.09	<b>0.60</b>	0.37
Pro-Social Punishment	0.06	-0.07	-0.04	-0.00	<b>0.48</b>	0.60
Dislike Having More	0.19	-0.01	<b>0.25</b>	-0.08	-0.22	0.62
Dislike Having Less	-0.08	-0.11	<b>0.40</b>	0.03	0.15	0.52
WTA	-0.05	- <b>0.51</b>	-0.06	0.10	0.00	0.31
Risk Aversion: Gains	0.03	<b>0.55</b>	0.11	0.06	-0.06	0.23
Risk Aversion: Losses	-0.11	<b>0.35</b>	-0.16	0.11	0.11	0.45
Risk Aversion: Gain / Loss	-0.02	<b>0.50</b>	-0.11	-0.02	-0.03	0.25
WTP	0.01	0.01	- <b>0.47</b>	0.12	0.06	0.46
Risk Aversion: CR Certain	-0.00	0.11	<b>0.53</b>	-0.00	0.04	0.39
Risk Aversion: CR Lottery	-0.03	-0.02	<b>0.42</b>	0.11	-0.03	0.57
Ambiguity Aversion	-0.05	-0.03	-0.04	<b>0.33</b>	- <b>0.29</b>	0.67
Compound Lottery Aversion	-0.14	-0.02	0.01	<b>0.39</b>	- <b>0.34</b>	0.50
Overestimation	0.07	-0.02	0.02	<b>0.45</b>	0.18	0.59
Overplacement	0.07	0.06	0.17	<b>0.38</b>	0.03	0.69
Overprecision	0.03	-0.01	-0.06	<b>0.56</b>	0.07	0.43
Patience	0.19	-0.18	-0.06	-0.05	- <b>0.28</b>	0.65
Percent of Variation	14%	13%	11%	8%	8%	47%

Notes: First five principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

Table E.4: Using 7 components, splits the Overconfidence component (from Table 7)

	Generosity		Risk Aversion:		Inequality		Uncertainty	Overconfidence		Unexplained
	WTA	WTP	Aversion/	WTP	Punishment	1		2		
Reciprocity: Low	<b>0.53</b>	0.01	-0.01	0.07	0.07	-0.01	-0.05	0.12	0.24	
Reciprocity: High	<b>0.55</b>	0.02	-0.02	0.08	0.08	0.04	-0.13	0.08	0.20	
Altruism	<b>0.31</b>	0.05	0.07	-0.01	-0.01	-0.05	0.38	-0.31	0.35	
Trust	<b>0.43</b>	0.01	-0.04	-0.07	-0.07	-0.06	0.27	-0.07	0.31	
Anti-Social Punishment	-0.01	0.00	0.03	<b>0.67</b>	<b>0.67</b>	-0.03	0.06	0.02	0.26	
Pro-Social Punishment	0.10	-0.04	-0.04	<b>0.60</b>	<b>0.60</b>	0.03	-0.04	-0.07	0.41	
Dislike Having More	0.18	0.00	<b>0.30</b>	-0.14	-0.14	0.15	0.07	<b>-0.29</b>	0.51	
Dislike Having Less	-0.03	-0.09	<b>0.38</b>	<b>0.25</b>	<b>0.25</b>	0.08	-0.14	0.14	0.45	
WTA	-0.03	<b>-0.50</b>	-0.06	0.04	0.04	0.11	0.04	0.02	0.31	
Risk Aversion: Gains	0.03	<b>0.55</b>	0.11	-0.05	-0.05	0.06	0.01	0.06	0.22	
Risk Aversion: Losses	-0.10	<b>0.36</b>	-0.16	0.15	0.15	0.10	0.04	0.01	0.42	
Risk Aversion: Gain / Loss	0.00	<b>0.50</b>	-0.11	-0.01	-0.01	0.03	-0.07	0.01	0.24	
WTP	-0.00	0.00	<b>-0.48</b>	0.01	0.01	-0.02	0.08	0.08	0.45	
Risk Aversion: CR Certain	-0.03	0.10	<b>0.52</b>	-0.01	-0.01	-0.11	0.02	0.13	0.38	
Risk Aversion: CR Lottery	-0.06	-0.03	<b>0.42</b>	-0.04	-0.04	0.03	0.14	0.05	0.55	
Ambiguity Aversion	0.07	0.01	-0.02	0.06	0.06	<b>0.71</b>	-0.07	0.01	0.24	
Compound Lottery Aversion	-0.08	0.01	0.02	-0.08	-0.08	<b>0.64</b>	0.10	-0.02	0.30	
Overestimation	-0.06	-0.05	0.01	0.04	0.04	0.00	<b>0.66</b>	0.05	0.36	
Overplacement	0.09	0.04	0.09	-0.03	-0.03	-0.02	0.01	<b>0.71</b>	0.31	
Overprecision	-0.03	-0.03	-0.12	-0.03	-0.03	0.07	<b>0.42</b>	<b>0.45</b>	0.33	
Patience	<b>0.25</b>	-0.18	-0.07	-0.23	-0.23	0.06	<b>-0.26</b>	0.19	0.54	
Percent of Variation	13%	13%	11%	8%	8%	7%	7%	6%	35%	

Notes: First seven principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

Table E.5: Using 8 components, as recommended by the Kaiser Criterion, splits Generosity and Overconfidence components (from Table 7)

	<i>Risk Aversion: WTA</i>	<i>Inequality Aversion/WTP</i>	<i>Reciprocity</i>	<i>Altruism / Trust / Dislike Having More</i>	<i>Punishment</i>	<i>Uncertainty</i>	<i>Overconfidence 1</i>	<i>Overconfidence 2</i>	<i>Unexplained</i>
Reciprocity: Low	0.00	0.05	<b>0.65</b>	0.00	0.01	0.00	0.06	0.02	0.11
Reciprocity: High	0.01	0.02	<b>0.64</b>	0.03	0.04	0.05	-0.06	0.01	0.12
Altruism	0.04	-0.05	-0.03	<b>0.66</b>	0.05	-0.06	0.09	-0.08	0.24
Trust	0.00	-0.13	0.18	<b>0.50</b>	-0.02	-0.07	0.05	0.09	0.25
Anti-Social Punishment	0.01	0.01	0.00	-0.02	<b>0.67</b>	-0.03	0.07	0.02	0.25
Pro-Social Punishment	-0.03	-0.06	0.07	0.05	<b>0.61</b>	0.03	-0.06	-0.05	0.40
Dislike Having More	-0.01	0.21	-0.07	<b>0.46</b>	-0.08	0.15	-0.15	-0.11	0.44
Dislike Having Less	-0.10	<b>0.33</b>	-0.13	0.05	<b>0.29</b>	0.07	-0.22	0.22	0.38
WTA	-0.50	-0.06	-0.01	-0.06	0.04	0.11	0.06	0.00	0.31
Risk Aversion: Gains	<b>0.55</b>	0.12	0.05	0.02	-0.05	0.06	0.02	0.05	0.22
Risk Aversion: Losses	<b>0.37</b>	-0.18	-0.15	0.02	0.17	0.09	0.00	0.05	0.39
Risk Aversion: Gain / Loss	<b>0.51</b>	-0.10	0.00	-0.02	-0.01	0.02	-0.07	0.01	0.24
WTP	0.02	-0.52	-0.09	0.04	0.04	-0.04	0.00	0.13	0.39
Risk Aversion: CR Certain	0.09	<b>0.56</b>	0.08	-0.09	-0.04	-0.09	0.12	0.06	0.31
Risk Aversion: CR Lottery	-0.04	<b>0.36</b>	-0.21	0.23	0.01	0.03	0.01	0.17	0.50
Ambiguity Aversion	0.01	-0.03	0.08	-0.02	0.06	<b>0.70</b>	-0.06	0.00	0.24
Compound Lottery Aversion	0.00	0.02	-0.06	-0.01	-0.09	<b>0.65</b>	0.12	-0.02	0.29
Overestimation	-0.05	0.04	0.02	0.07	-0.01	0.02	<b>0.73</b>	-0.01	0.25
Overplacement	0.04	0.04	0.03	-0.06	0.00	-0.04	-0.05	<b>0.75</b>	0.25
Overprecision	-0.03	-0.14	-0.05	0.02	-0.02	0.07	<b>0.40</b>	<b>0.46</b>	0.33
Patience	-0.18	-0.14	0.09	0.13	-0.18	0.04	-0.40	<b>0.29</b>	0.44
Percent of Variation	13%	10%	10%	8%	8%	7%	7%	6%	30%

Notes: First eight principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

Table E.6: The patterns in Table 7 are largely robust to the omission of sample weights.

	<i>Generosity</i>	<i>Risk Aversion: WTA</i>	<i>Inequality Aversion: WTP</i>	<i>Impulsivity</i>	<i>Overconfidence</i>	<i>Uncertainty</i>	Unexplained
Reciprocity: Low	<b>0.55</b>	0.01	0.03	0.12	-0.10	0.04	0.23
Reciprocity: High	<b>0.55</b>	0.00	-0.01	0.06	-0.12	0.08	0.21
Altruism	<b>0.36</b>	0.03	0.02	-0.08	0.24	-0.13	0.51
Trust	<b>0.44</b>	-0.01	-0.06	-0.12	0.22	-0.11	0.32
Anti-Social Punishment	0.04	-0.01	0.05	<b>0.61</b>	0.04	-0.01	0.34
Pro-Social Punishment	0.10	-0.06	-0.04	<b>0.50</b>	-0.07	0.06	0.57
Dislike Having More	0.17	0.02	<b>0.30</b>	-0.27	0.02	0.07	0.59
Dislike Having Less	-0.12	-0.12	<b>0.34</b>	0.10	0.05	0.01	0.59
WTA	-0.01	<b>-0.51</b>	-0.08	0.05	-0.05	-0.03	0.37
Risk Aversion: Gains	0.02	<b>0.55</b>	0.12	-0.03	0.07	0.02	0.26
Risk Aversion: Losses	-0.03	<b>0.36</b>	-0.21	0.11	-0.05	0.00	0.43
Risk Aversion: Gain / Loss	-0.02	<b>0.48</b>	-0.13	0.00	-0.07	-0.01	0.27
WTP	0.00	0.00	<b>-0.46</b>	0.02	0.08	-0.03	0.51
Risk Aversion: CR Certain	0.01	0.12	<b>0.53</b>	0.07	-0.04	-0.06	0.42
Risk Aversion: CR Lottery	-0.02	-0.04	<b>0.44</b>	0.01	0.05	0.01	0.54
Ambiguity Aversion	0.03	0.00	-0.03	0.00	-0.01	<b>0.69</b>	0.27
Compound Lottery Aversion	-0.01	0.02	0.02	0.00	0.05	<b>0.67</b>	0.29
Overestimation	-0.02	-0.07	0.02	<b>0.29</b>	<b>0.36</b>	0.00	0.57
Overplacement	-0.04	0.04	0.04	-0.13	<b>0.60</b>	0.04	0.46
Overprecision	-0.04	0.00	-0.08	0.12	<b>0.59</b>	0.03	0.36
Patience	0.08	-0.20	-0.09	- <b>0.35</b>	0.04	0.09	0.67
Percent of Variation	12%	12%	11%	8%	7%	7%	42%

Notes: First six principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

Table E.7: The patterns in Table 7 are largely robust to using the average of measures, rather than ORIV.

	<i>Generosity</i>	<i>Risk Aversion: WTA</i>	<i>Inequality Aversion: WTP</i>	<i>Impulsivity</i>	<i>Uncertainty</i>	<i>Overconfidence</i>	<i>Unexplained</i>
Reciprocity: Low	<b>0.49</b>	-0.01	0.00	0.03	-0.03	0.02	0.36
Reciprocity: High	<b>0.52</b>	0.01	-0.01	0.03	0.01	-0.05	0.29
Altruism	<b>0.40</b>	0.06	0.03	0.07	0.00	0.03	0.56
Trust	<b>0.48</b>	0.01	-0.06	-0.01	-0.02	0.08	0.38
Anti-Social Punishment	-0.02	0.00	0.03	<b>0.68</b>	-0.01	0.01	0.27
Pro-Social Punishment	0.08	-0.04	-0.03	<b>0.62</b>	0.02	-0.06	0.40
Dislike Having More	0.20	0.00	<b>0.27</b>	-0.12	0.13	-0.17	0.63
Dislike Having Less	-0.06	-0.10	<b>0.39</b>	0.19	0.05	0.00	0.56
WTA	-0.04	<b>-0.50</b>	-0.06	0.04	0.09	0.05	0.39
Risk Aversion: Gains	0.04	<b>0.55</b>	0.11	-0.04	0.05	0.03	0.32
Risk Aversion: Losses	-0.09	<b>0.36</b>	-0.16	0.13	0.07	0.06	0.51
Risk Aversion: Gain / Loss	-0.01	<b>0.50</b>	-0.10	-0.02	0.02	-0.04	0.33
WTP	0.00	0.00	<b>-0.48</b>	0.02	-0.02	0.13	0.49
Risk Aversion: CR Certain	-0.02	0.10	<b>0.53</b>	-0.01	-0.09	0.07	0.43
Risk Aversion: CR Lottery	-0.03	-0.03	<b>0.42</b>	-0.01	0.04	0.11	0.61
Ambiguity Aversion	0.04	0.00	-0.02	0.04	<b>0.70</b>	-0.03	0.24
Compound Lottery Aversion	-0.05	0.01	0.01	-0.04	<b>0.67</b>	0.05	0.27
Overestimation	0.03	-0.03	0.00	0.08	0.01	<b>0.43</b>	0.72
Overplacement	0.03	0.03	0.12	-0.06	-0.04	<b>0.55</b>	0.54
Overprecision	0.00	-0.02	-0.06	-0.02	0.04	<b>0.65</b>	0.37
Patience	0.19	-0.20	-0.05	- <b>0.25</b>	0.04	-0.05	0.69
Percent of Variation	13%	11%	9%	7%	7%	7%	44%

Notes: First six principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.



Table E.8: The patterns in Table 7 are largely robust to using Spearman rank-order correlations.

	<i>Generosity</i>	<i>Risk Aversion:</i> <i>WTA</i>	<i>Inequality Aversion:</i> <i>WTP</i>	<i>Impulsivity</i>	<i>Uncertainty</i>	<i>Overconfidence</i>	Unexplained
Reciprocity: Low	<b>0.49</b>	-0.01	-0.01	0.02	0.04	-0.06	0.39
Reciprocity: High	<b>0.52</b>	-0.02	-0.05	-0.03	0.06	-0.07	0.31
Altruism	<b>0.42</b>	0.05	0.06	0.11	-0.05	0.09	0.53
Trust	<b>0.48</b>	0.00	-0.01	0.02	-0.09	0.10	0.39
Anti-Social Punishment	-0.02	0.00	0.05	<b>0.67</b>	-0.01	0.03	0.31
Pro-Social Punishment	0.08	-0.03	-0.04	<b>0.61</b>	0.03	-0.06	0.45
Dislike Having More	0.20	0.06	<b>0.30</b>	-0.21	0.06	-0.06	0.63
Dislike Having Less	-0.10	-0.10	<b>0.32</b>	0.14	0.07	0.05	0.68
WTA	-0.02	<b>-0.51</b>	-0.13	0.00	-0.02	-0.04	0.46
Risk Aversion: Gains	0.02	<b>0.54</b>	0.11	-0.05	0.01	0.05	0.38
Risk Aversion: Losses	-0.01	<b>0.35</b>	<b>-0.27</b>	0.05	0.00	-0.04	0.49
Risk Aversion: Gain / Loss	-0.03	<b>0.48</b>	-0.15	-0.01	0.00	-0.05	0.38
WTP	0.04	0.00	<b>-0.48</b>	0.00	0.01	0.04	0.53
Risk Aversion: CR Certain	-0.02	0.14	<b>0.48</b>	0.01	-0.08	-0.03	0.54
Risk Aversion: CR Lottery	0.01	-0.01	<b>0.44</b>	0.06	0.02	0.00	0.61
Ambiguity Aversion	0.00	0.00	-0.04	-0.03	<b>0.70</b>	0.00	0.28
Compound Lottery Aversion	0.00	0.02	0.03	0.04	<b>0.69</b>	0.01	0.30
Overestimation	-0.02	-0.06	0.03	0.11	0.02	<b>0.41</b>	0.72
Overplacement	0.00	0.04	0.03	-0.09	0.03	<b>0.59</b>	0.52
Overprecision	0.01	0.01	-0.05	0.02	-0.02	<b>0.65</b>	0.40
Patience	0.10	-0.23	0.05	<b>-0.27</b>	0.04	0.05	0.73
Percent of Variation	12%	10%	9%	7%	7%	7%	48%

Notes: First six principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

Table E.9: The patterns in Table 7 are largely robust to using Factor Analysis.

	<i>Generosity</i>	<i>Risk Aversion: WTA</i>	<i>Inequality Aversion: WTP</i>	<i>Uncertainty</i>	<i>Impulsivity</i>	<i>Overconfidence</i>	<i>Uniqueness</i>
Reciprocity: Low	<b>0.80</b>	-0.02	0.02	-0.02	0.03	-0.04	0.35
Reciprocity: High	<b>0.85</b>	0.00	0.00	0.01	0.01	-0.08	0.27
Altruism	<b>0.49</b>	0.05	0.04	-0.09	-0.03	<b>0.33</b>	0.63
Trust	<b>0.62</b>	0.01	-0.06	-0.09	-0.09	<b>0.35</b>	0.47
Anti-Social Punishment	-0.02	0.01	0.02	-0.03	<b>0.63</b>	0.02	0.60
Pro-Social Punishment	0.10	-0.03	-0.03	-0.01	<b>0.50</b>	-0.01	0.74
Dislike Having More	<b>0.25</b>	-0.09	<b>0.31</b>	0.05	-0.18	-0.02	0.80
Dislike Having Less	-0.07	-0.22	<b>0.44</b>	0.04	0.17	-0.02	0.72
WTA	-0.04	- <b>0.65</b>	-0.03	0.12	0.04	0.01	0.56
Risk Aversion: Gains	0.02	<b>0.69</b>	0.08	0.03	-0.02	0.04	0.52
Risk Aversion: Losses	-0.15	<b>0.52</b>	- <b>0.25</b>	0.07	0.17	0.05	0.61
Risk Aversion: Gain / Loss	-0.02	<b>0.72</b>	-0.19	0.00	0.00	-0.04	0.44
WTP	-0.03	0.12	- <b>0.54</b>	0.00	0.05	0.12	0.67
Risk Aversion: CR Certain	-0.01	0.00	<b>0.58</b>	-0.07	0.01	0.01	0.66
Risk Aversion: CR Lottery	-0.06	-0.14	<b>0.44</b>	0.03	-0.01	0.13	0.76
Ambiguity Aversion	0.01	-0.02	-0.03	<b>0.63</b>	0.01	-0.02	0.60
Compound Lottery Aversion	-0.12	-0.02	0.00	<b>0.62</b>	-0.05	0.04	0.60
Overestimation	-0.01	-0.04	-0.02	0.05	0.11	0.24	0.92
Overplacement	-0.02	-0.01	0.08	0.03	0.04	<b>0.31</b>	0.89
Overprecision	-0.08	-0.01	-0.11	0.11	0.08	<b>0.42</b>	0.79
Patience	0.23	-0.23	-0.01	0.02	-0.24	0.01	0.84
Proportion of Variance	<b>0.33</b>	0.28	0.19	0.13	0.13	0.09	

Notes: First six factors, rotated. Factor loadings greater than or equal to 0.25 in bold.

Table E.10: Adding an additional risk aversion measure switches the ranking of several components, while leave the structure largely unchanged.

	Risk Aversion: WTA	Generosity	Inequality Aversion/ WTP	Uncertainty	Overconfidence	Impulsivity	Unexplained
Reciprocity: Low	0.00	<b>0.50</b>	0.00	-0.02	0.02	0.05	0.29
Reciprocity: High	0.01	<b>0.51</b>	-0.02	0.04	-0.07	0.06	0.27
Altruism	0.05	<b>0.39</b>	0.03	-0.02	0.04	0.05	0.55
Trust	0.01	<b>0.47</b>	-0.06	-0.05	0.12	-0.04	0.33
Anti-social Punishment	0.00	-0.01	0.03	-0.02	0.05	<b>0.67</b>	0.26
Pro-social Punishment	-0.04	0.09	-0.04	0.02	-0.08	<b>0.59</b>	0.42
Dislike Having More	0.02	0.22	<b>0.27</b>	0.14	-0.18	-0.11	0.58
Dislike Having Less	-0.08	-0.06	<b>0.40</b>	0.05	-0.02	0.21	0.49
WTA	<b>-0.46</b>	-0.04	-0.05	0.06	0.06	0.04	0.30
Risk Aversion: Gains	<b>0.52</b>	0.04	0.10	0.09	0.04	-0.05	0.20
Risk Aversion: Losses	<b>0.32</b>	-0.08	-0.18	0.16	0.04	0.16	0.44
Risk Aversion: Gain/Loss	<b>0.45</b>	0.00	-0.12	0.10	-0.05	-0.01	0.27
<b>Risk Aversion: Urn</b>	<b>0.40</b>	-0.04	0.03	<b>-0.35</b>	0.01	0.00	0.18
WTP	0.00	0.00	<b>-0.47</b>	-0.01	0.14	0.01	0.45
Risk Aversion: CR Certain	0.09	-0.02	<b>0.52</b>	-0.07	0.07	-0.01	0.39
Risk Aversion: CR Lottery	-0.03	-0.03	<b>0.42</b>	0.05	0.11	-0.02	0.57
Ambiguity Aversion	0.03	0.04	-0.01	<b>0.65</b>	-0.04	0.05	0.29
Compound Lottery Aversion	0.03	-0.07	0.02	<b>0.61</b>	0.08	-0.07	0.30
Overestimation	-0.03	0.04	0.01	0.00	<b>0.51</b>	0.10	0.54
Overplacement	0.03	0.03	0.15	-0.03	<b>0.49</b>	-0.09	0.57
Overprecision	-0.01	-0.01	-0.07	0.04	<b>0.62</b>	-0.02	0.33
Patience	-0.17	0.19	-0.05	0.02	-0.05	<b>-0.28</b>	0.65
Percent of Variation	14%	13%	10%	8%	8%	7%	39%

Notes: First six principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.

Table E.11: Removing “Dislike Having More” leaves the overall structure of components unchanged.

	Generosity	Risk Aversion: WTA	Inequality Aversion/ WTP	Overconfidence	Impulsivity	Uncertainty	Unexplained
Negative Reciprocity	<b>0.52</b>	0.01	0.03	-0.01	0.04	0.03	0.26
Positive Reciprocity	<b>0.53</b>	0.02	0.01	-0.09	0.05	0.07	0.24
Altruism	<b>0.39</b>	0.05	0.02	0.06	0.05	-0.07	0.56
Trust	<b>0.47</b>	0.01	-0.06	0.13	-0.04	-0.07	0.34
Anti-Social Punishment	-0.01	0.00	0.03	0.05	<b>0.67</b>	-0.02	0.27
Pro-Social Punishment	0.08	-0.05	-0.05	-0.07	<b>0.60</b>	0.03	0.41
Dislike Having Less	-0.07	-0.11	<b>0.38</b>	-0.01	0.22	0.05	0.51
WTA	-0.04	<b>-0.50</b>	-0.07	0.06	0.04	0.10	<b>0.31</b>
Risk Aversion: Gains	0.04	<b>0.55</b>	0.12	0.03	-0.05	0.06	0.23
Risk Aversion: Losses	-0.11	<b>0.36</b>	-0.18	0.08	0.17	0.06	0.41
Risk Aversion: Gain/Loss	-0.01	<b>0.50</b>	-0.10	-0.04	-0.01	0.03	0.24
WTP	-0.03	0.01	<b>-0.49</b>	0.16	0.02	-0.05	0.44
Risk Aversion: CR Certain	0.00	0.11	<b>0.56</b>	0.03	-0.02	-0.07	0.35
Risk Aversion: CR Lottery	-0.01	-0.03	<b>0.44</b>	0.09	-0.03	0.05	0.55
Ambiguity Aversion	0.06	0.00	-0.01	-0.04	0.05	<b>0.72</b>	0.23
Compound Lottery Aversion	-0.06	0.01	0.02	0.08	-0.07	<b>0.65</b>	0.30
Overestimation	0.02	-0.03	0.00	<b>0.52</b>	0.11	-0.02	0.52
Overplacement	0.03	0.04	0.17	<b>0.48</b>	-0.09	-0.04	0.57
Overprecision	-0.02	-0.02	-0.08	<b>0.63</b>	-0.02	0.05	0.32
Patience	0.20	-0.18	-0.04	-0.06	<b>-0.28</b>	0.06	0.64
Percent of Variation	14%	13%	10%	8%	8%	7%	39%

Notes: First six principal components using the varimax rotation. Weights greater than or equal to 0.25 in bold.